

# MA/STAT 538, Spring 2009, (ip)

1. Recall the definition that a random variable  $X$  is called *integrable* if  $E|X| < \infty$ .

- (a) Prove that if  $X$  and  $Y$  are two integrable and *independent* random variables, then  $XY$  is also integrable and that  $E(XY) = (EX)(EY)$ .
- (b) Give an example of two integrable random variables  $X$  and  $Y$  such that  $XY$  is not integrable.
- (c) Give an example of two random variables  $X$  and  $Y$  such that  $X + Y$  is integrable and yet  $X$  and  $Y$  are not integrable.

$$(a) \quad E|XY| = \text{(LMCT)} \lim_{N \rightarrow \infty} E|X| \mathbb{1}_{\{|X| \leq N\}} |Y| \mathbb{1}_{\{|Y| \leq N\}}$$

$$\begin{aligned} & \text{X, Y-ind.} \\ & = \lim_{N \rightarrow \infty} \left( E|X| \mathbb{1}_{\{|X| \leq N\}} \right) \left( E|Y| \mathbb{1}_{\{|Y| \leq N\}} \right) \end{aligned}$$

$$\text{LMCT} = E|X| E|Y| < \infty$$

The same proof gives

$$E(XY) = EX EY$$

$$(b) \quad (\Omega = [0, 1], \mathcal{B}, P = x)$$

$$\text{Let } X=Y, \quad X(x) = \frac{1}{\sqrt{x}} \text{ — integrable}$$

$$Y(x) = \frac{1}{\sqrt{x}} \text{ — integrable}$$

$$\text{but } XY = \frac{1}{x} \text{ — not integrable.}$$

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$$\left( \Omega = [0, 1], \mathcal{B}, \lambda \right)$$

$$(c) \quad X = \frac{1}{x} \quad (\text{not integrable})$$

$$Y = -\frac{1}{x} \quad (\text{not integrable})$$

$$\text{but } X + Y \equiv 0 \quad (\text{integrable}).$$

2. Let  $\{B_n\}_{n \geq 1}$  be a sequence of events. Suppose there is a  $\delta > 0$  such that  $P(B_n) \geq \delta$  for all  $n$ . Prove that  $P(B_n \text{ i.o.}) \geq \delta$ .

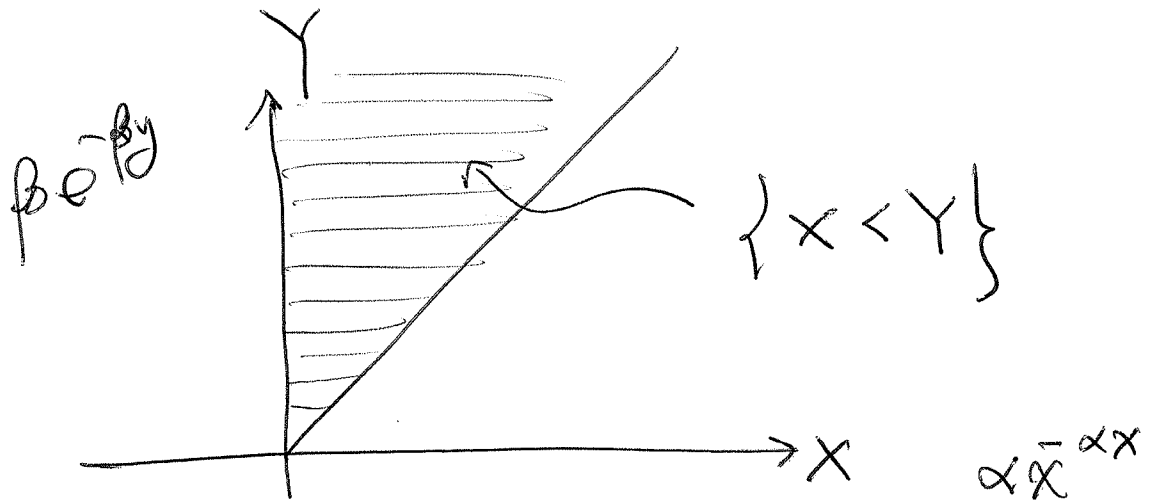
M1

$$P(\overline{\lim B_n}) \geq \overline{\lim} P(B_n) \geq \delta > 0$$

M2

$$\begin{aligned} P(\overline{\lim B_n}) &= P\left(\bigcap_N \bigcup_{n=N}^{\infty} B_n\right) \\ &= \lim_{N \rightarrow \infty} P\left(\bigcup_{n=N}^{\infty} B_n\right) \\ &\geq \lim_{N \rightarrow \infty} P(B_N) \geq \delta > 0 \end{aligned}$$

3. Let  $X$  and  $Y$  be two independent, exponentially distributed random variables with parameter  $\alpha$  and  $\beta$ , i.e. the pdfs of  $X$  and  $Y$  are given by  $\alpha e^{-\alpha x}$  and  $\beta e^{-\beta y}$  for  $x, y \geq 0$ . Find  $P(X < Y)$ .



$$P(X < Y) = \int_0^{\infty} \left[ \int_0^y \alpha e^{-\alpha x} dx \right] \beta e^{-\beta y} dy$$

$$= \int_0^{\infty} (1 - e^{-\alpha y}) \beta e^{-\beta y} dy$$

$$= \int_0^{\infty} \beta e^{-\beta y} dy - \beta \int_0^{\infty} e^{-(\alpha + \beta)y} dy$$

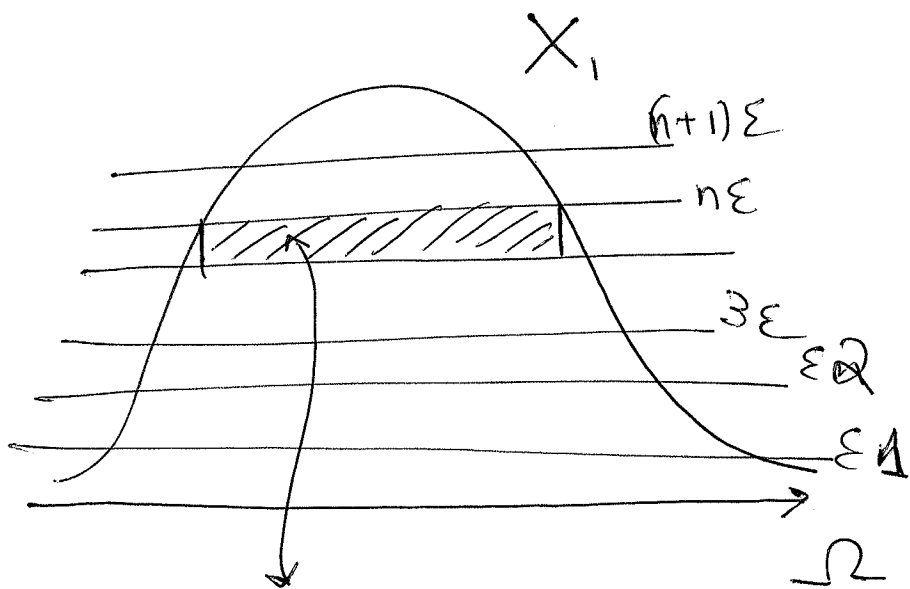
$$= 1 - \frac{\beta}{\alpha + \beta} = \boxed{\frac{\alpha}{\alpha + \beta}}$$

4. Let  $X_1, X_2, \dots$  be i.i.d. non-negative random variables. Prove that  $\limsup_n \frac{X_n}{n}$  equals zero or infinity depending on whether  $EX_1$  is finite or infinity.

$\mathbb{P} \quad EX_1 < \infty$

then  $\forall \varepsilon$ , consider  $\sum_{n=1}^{\infty} P(X_n \geq \varepsilon n)$

$$= \sum_{n=1}^{\infty} P(X_1 \geq \varepsilon n)$$



$$\text{Area} = \varepsilon P(X_1 \geq n\varepsilon)$$

$$= \sum_{n=1}^{\infty} P(X_1 \geq n\varepsilon) \varepsilon \leq \frac{1}{\varepsilon} EX_1 < \infty$$

$$\text{B-C.I} \Rightarrow P(X_n \geq \varepsilon n \text{ i.o.}) = 0$$

$$\text{i.e. } \forall \varepsilon, P\left(\frac{X_n}{n} \geq \varepsilon \text{ i.o.}\right) = 0 \Rightarrow \lim_{n \rightarrow \infty} \frac{X_n}{n} \leq \varepsilon \text{ P.a.s. } (\varepsilon \rightarrow 0)$$

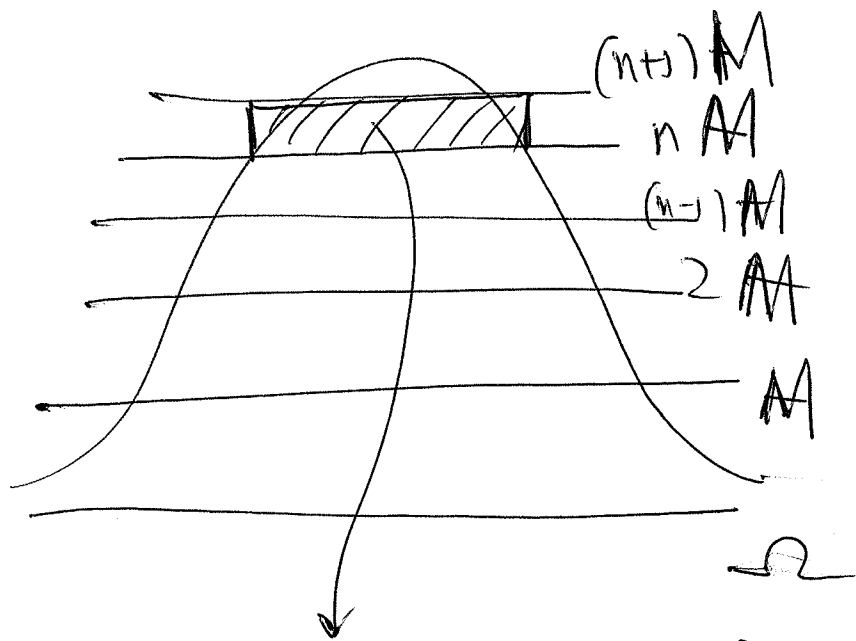
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$X_n - \text{i.i.d}$

Similarly, if  $EX_1 = \infty$

then  $\forall M$ , Consider  $\sum_{n=1}^{\infty} P(X_n \geq Mn)$

$$= \sum_{n=1}^{\infty} P(X_1 \geq Mn)$$



$$\text{Area} = P(X \geq nM)$$

$$= \frac{1}{M} \sum_{n=1}^{\infty} P(X_1 \geq Mn) M$$

$$\geq \frac{1}{M} EX_1 = \infty$$

B-C-II  $\Rightarrow P(X_n \geq Mn \text{ i.o.}) = 1$   
 (ind)  $X_n$ 's i.e.  $\forall M, P(\frac{X_n}{n} \geq M \text{ i.o.}) = 1 \Rightarrow \lim_{M \rightarrow \infty} \frac{X_n}{n} \geq M$

5. Let  $X_1, X_2, \dots$  be a sequence of i.i.d., integrable random variables. Let  $T$  be another random variable which takes value in positive integers and is independent of all of the  $X_i$ 's. Define

$$S_T = \sum_{i=1}^T X_i$$

Prove the following:

- (a)  $ES_T = E(T)E(X_1)$ .  
 (b)  $\text{Var}(S_T) = [E(X_1)]^2 \text{Var}(T) + E(T)\text{Var}(X_1)$ .

$$(a) \quad ES_T = E[X_1 + X_2 + \dots + X_T]$$

$$= \sum_{n=1}^{\infty} E(X_1 + X_2 + \dots + X_n) \mathbb{1}_{\{T=n\}}$$

$$= \sum_{n=1}^{\infty} E(X_1 + \dots + X_n) E \mathbb{1}_{\{T=n\}}$$

$$= \sum_{n=1}^{\infty} n(E X_1) P(T=n)$$

$$= EX_1 \sum_{n=1}^{\infty} n P(T=n) = EX ET$$

$$\text{Var}(S_T) = E(S_T^2) - (E S_T)^2$$

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$$(b) \quad E(S_T^2)$$

$$= \sum_{n=1}^{\infty} E(X_1 + \dots + X_n)^2 P(T=n) \quad \begin{matrix} (a \ b \ c)^2 \\ a^2 + b^2 + c^2 \\ 2ab + 2bc + 2ca \end{matrix}$$

$$= \sum_{n=1}^{\infty} \left( n E(X_i^2) + \underbrace{n(n-1)}_{\substack{3 \cdot 2 \\ a \ b \ c \ d}} E X_i E X_j \right) P(T=n) \quad \begin{matrix} 2ab \ 2bc \ 2cd \\ 2ac \ 2ad \\ 2bd \\ 4 \times 3 \end{matrix}$$

$$= \sum_{n=1}^{\infty} \left[ n E(X_i^2) + n(n-1) (E X_i)^2 \right] P(T=n)$$

$$= \sum_{n=1}^{\infty} n P(T=n) E(X_i^2) + (E X_i)^2 \sum_{n=1}^{\infty} (n^2 - n) P(T=n)$$

$$= \cancel{E(X_i^2) E(T)} + (E X_i)^2 \left[ \cancel{E(T^2)} - \cancel{(E T)^2} \right]$$

$$= \left( E(X_i^2) - (E X_i)^2 \right) E T + (E X_i)^2 E(T^2)$$

6. In this problem,  $\{N_n\}_{n \geq 1}$  denotes some sequence of positive-integer-valued random variables such that  $N_n \rightarrow \infty$  with probability one (i.e. P a.s.).

(a) Let  $\{X_n\}_{n \geq 1}$  be a sequence of random variables converging to 0 with probability one. Prove that  $X_{N_n}$  converges to zero with probability one.

(b) Give an example of a  $\{N_n\}$  (satisfying the stated assumption at the beginning of this problem) and a sequence of random variables  $\{X_n\}$  which converges to zero in probability but  $X_{N_n}$  does not converge to zero in probability (and hence of course, cannot converge to zero with probability one).

(c) Give an example of a sequence of random variables  $\{X_n\}$  which converges to zero in probability but  $\frac{X_1 + X_2 + \dots + X_n}{n}$  does not converge to zero in probability.

$$(a) \quad X_n \rightarrow 0 \text{ P a.s.}$$

$$\text{i.e. } \exists \Omega_1 \text{ s.t. } P(\Omega_1) = 1 \text{ \& } \forall \omega \in \Omega_1, \quad X_n(\omega) \xrightarrow{n \rightarrow \infty} 0$$

(fixed)

$$N_n \rightarrow \infty \text{ P a.s.}$$

$$\text{i.e. } \exists \Omega_2 \text{ s.t. } P(\Omega_2) = 1 \text{ \& } \forall \omega \in \Omega_2, \quad N_n(\omega) \xrightarrow{n \rightarrow \infty} \infty$$

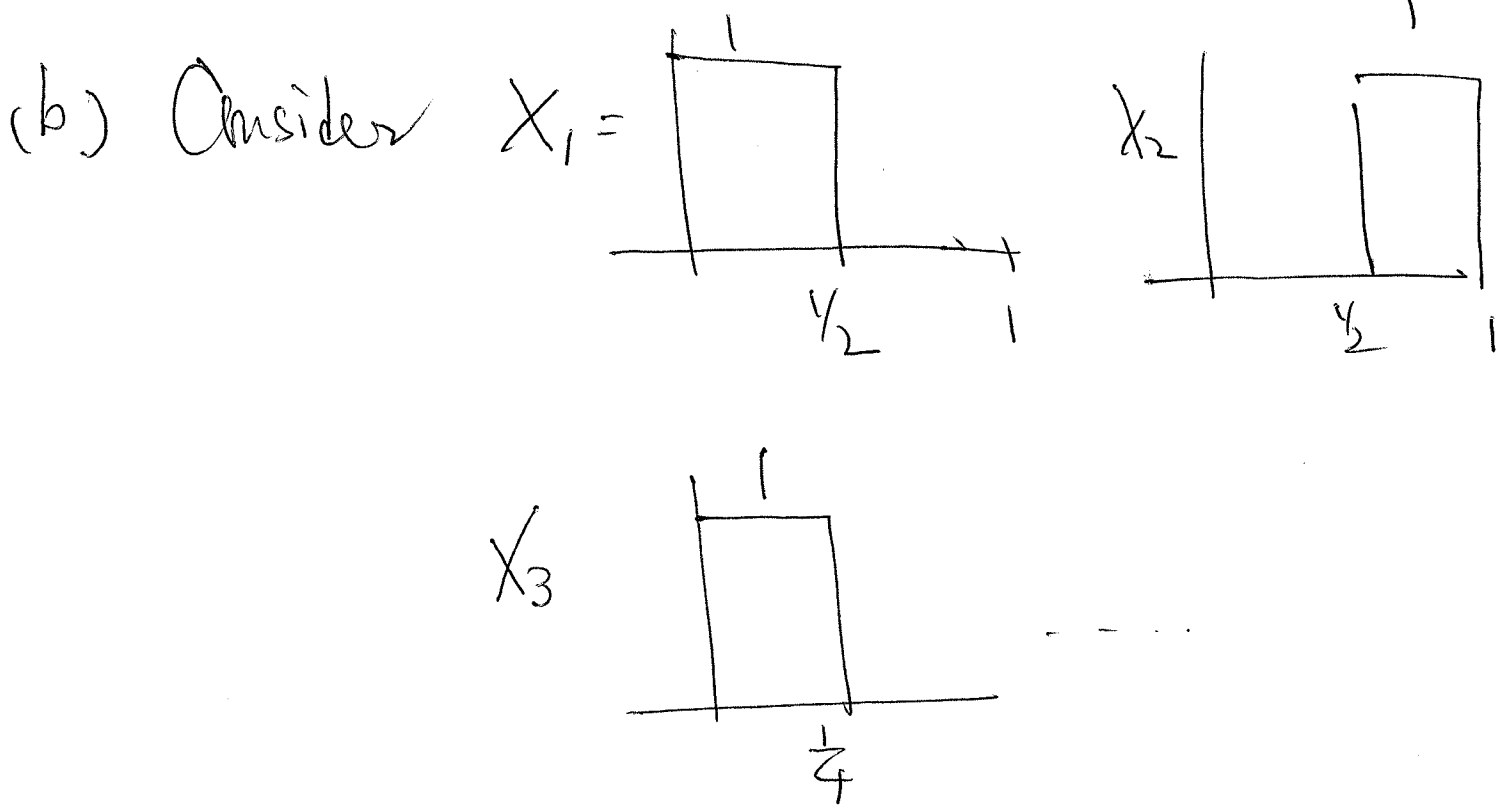
(fixed)

Hence  $\forall \omega \in \Omega_1 \cap \Omega_2$  (fixed)      Note  $P(\Omega_1 \cap \Omega_2) = 1$

$$\underbrace{X_{N_n(\omega)}(\omega)}_{\infty} \xrightarrow{n \rightarrow \infty} 0$$

fixed

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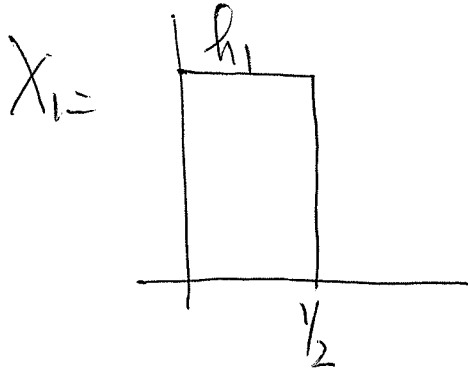
$\forall \omega \in (0, 1)$ , let  $N_n(\omega) =$  the  $n^{\text{th}}$  index  $k$  s.t.  $X_k(\omega) = 1$

then  $N_n(\omega) \xrightarrow{n \rightarrow \infty} \infty \quad \forall \omega \text{ (fixed)}$

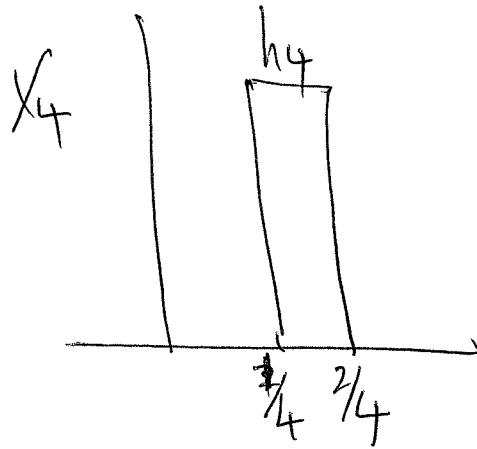
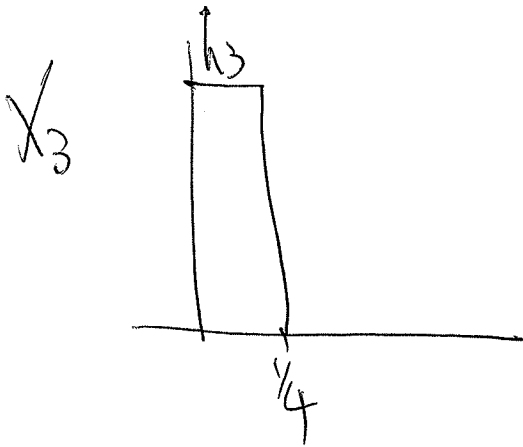
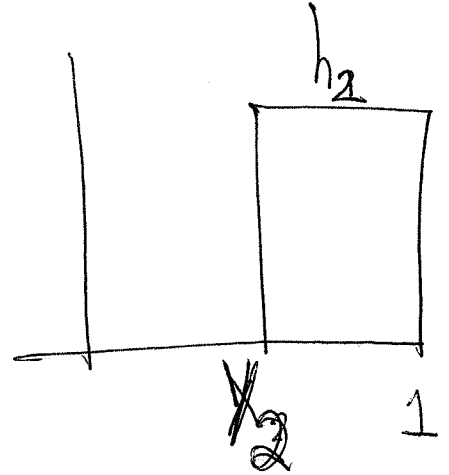
but  $X_{N_n(\omega)}(\omega) \equiv 1 \quad \forall n, \omega !$

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(c)



$X_2$



eg  $h_n = 2^{n+2}$

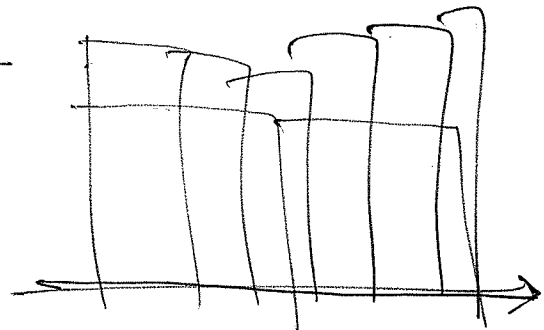
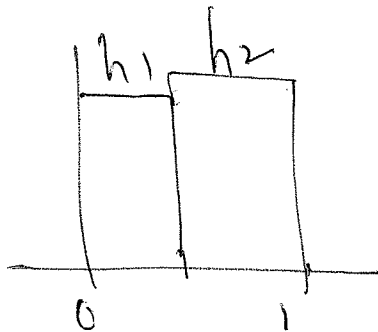
Choose  $h_1, h_2, \dots$  SO LARGE that

~~XXXXXXXXXX~~

$$\underbrace{X_1(\omega) + X_2(\omega) + \dots + X_n(\omega)}_n \geq \frac{1}{2} \text{ (for example)} \neq 0. \quad \forall n$$

eg.

$$\frac{X_1 + X_2}{2} =$$



eg.

$$\frac{X_1 + X_2 + X_3 + X_4 + X_5 + X_6}{6}$$