

Claim If $\{X_n; n \geq 1\}$ sequence
of iid Cauchy r.v. Then

$$\bar{X}_n \xrightarrow{(d)} z$$

But

$$\bar{X}_n \not\xrightarrow{D}$$

$\forall n, X_n \sim \text{Cauchy}$

Density

$$f(x) = \frac{1}{\pi(1+x^2)}$$

Cauchy random variable (1)

Notation:

Cauchy(α), with $\alpha \in \mathbb{R}$

State space:

\mathbb{R}

Density:

$$f(x) = \frac{1}{\pi} \frac{1}{1 + (x - \alpha)^2}$$

Expected value and variance:

Not defined (divergent integrals)!

Fact If $X \sim \text{Cauchy}$, $X \notin L^1$:

$$E[X] = \frac{1}{\pi} \int_{\mathbb{R}} \frac{x}{1+x^2} dx \rightarrow \text{not } 0 \text{ divergent}$$

Recall $X \in L^1(\Omega) \Leftrightarrow E[|X|] < \infty$
Here

$$E[|X|] = \frac{2}{\pi} \int_0^{\infty} \frac{x}{1+x^2} dx = \infty$$

$g(x) \sim \frac{1}{x}$

Riemann's rule. If at ∞ we have

$g(x) \sim \frac{1}{x^\alpha}$, then $\int_0^{\infty} g(x) dx < \infty$
iff $\alpha > 1$

$$\lim_{x \rightarrow \infty} \frac{g(x)}{x^{-\alpha}} = 1$$

Cauchy random variable (2)

Use 1: Trigonometric function of a uniform r.v

Namely if

- $X \sim \mathcal{U}([-\frac{\pi}{2}, \frac{\pi}{2}])$
- $Y = \tan(X)$

check at home

Then $Y \sim \text{Cauchy} \equiv \text{Cauchy}(0)$

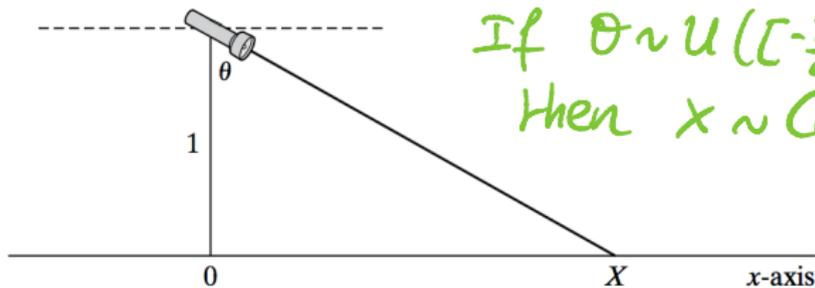
Use 2:

Typical example of r.v with no mean *(not in $L^1(\mathbb{R})$)*

Example: beam (1)

Experiment:

- Narrow-beam flashlight spun around its center
- Center located a unit distance from the x -axis
- X = point at which the beam intersects the x -axis when the flashlight has stopped spinning



$x = \tan(\theta)$
If $\theta \sim U\left(-\frac{\pi}{2}, \frac{\pi}{2}\right)$
Then $x \sim \text{Cauchy}$

Characteristic function of $X \sim \text{Cauchy}$

Define $\phi(t) = \mathbb{E}[e^{itx}]$

Here

$$\phi(t) = \int_{\mathbb{R}} \frac{e^{itx}}{1+x^2} dx$$

$= \cos(tx) + i \sin(tx)$

Method to compute ϕ

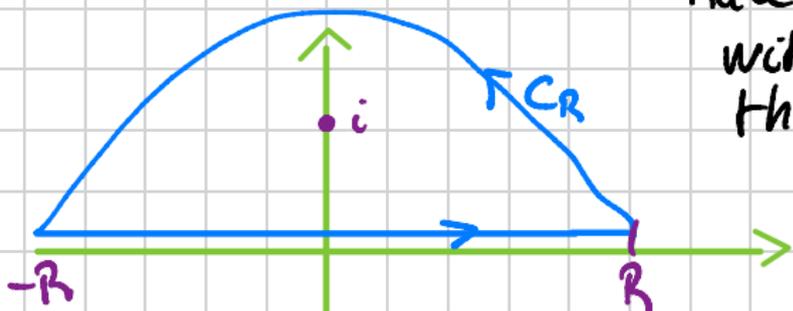
↳ Residues

$$\phi(t) = \int_{\mathbb{R}} \frac{e^{itx}}{\pi(1+x^2)} dx$$

Set $g_t(z) = \frac{e^{itz}}{\pi(1+z^2)}$

$g_t(z)$ has 2 poles: $z = \pm i$

Natural contour:



Rule: If $|g_t(z)| \leq \frac{C}{|z|^p}$
with $p > 1$,
then

$$\lim_{R \rightarrow \infty} \int_{C_R} g_t(z) dz = 0$$

If $t > 0$, $|g_t(z)| \leq \frac{1}{|1+z^2|} \leq \frac{e}{|z|^2}$
 $\Rightarrow \lim_{R \rightarrow \infty} \int_{C_R} g_t(z) dz = 0$

Thus $\phi(t) = \int_{-\infty}^{\infty} g_t(x) dx$ for $t > 0$

$$\lim_{R \rightarrow \infty} \int_{-R}^R g_t(x) dx = 2\pi i \operatorname{Res}(g_t, i)$$

$$= 2\pi i \lim_{z \rightarrow i} (z-i) g_t(z) = e^{-t}$$

$$g_t(z) = \frac{e^{itz}}{\pi(1+z^2)} = \frac{e^{itz}}{\pi(z-i)(z+i)}$$

$$\lim_{z \rightarrow i} (z-i) g_t(z) = \lim_{z \rightarrow i} \frac{e^{itz}}{\pi(z+i)}$$

$$= \frac{e^{-t}}{2i\pi}$$

For $t < 0$, we get

$$\phi(t) = e^t$$

Conclusion : if $X \sim \text{Cauchy}$

$$\phi(t) = e^{-|t|}$$

Example: beam (2)

Model:

- We assume $\theta \sim \mathcal{U}([-\frac{\pi}{2}, \frac{\pi}{2}])$
- We have $X \sim \tan(\theta)$

Conclusion:

$X \sim \text{Cauchy}$

Example with no WLLN

Rmk: This an example of $X_i \notin L^1$, for which LLN is far from being satisfied

Proposition 19.

We consider

- $\{X_n; n \geq 1\}$ sequence of i.i.d random variables
- $S_n = \sum_{i=1}^n X_i$ and $\bar{X}_n = \frac{1}{n}S_n$
- $X_1 \sim \text{Cauchy}$

Then

$\bar{X}_n \xrightarrow{(d)} \text{Cauchy}$, but \bar{X}_n does not converge in P

Char. function of \bar{X}_n

$$\phi_n(t) = \mathbb{E}[e^{it\bar{X}_n}] = \mathbb{E}[e^{i\frac{t}{n}\sum_{i=1}^n X_i}]$$

$$\stackrel{II}{=} \prod_{i=1}^n \mathbb{E}[e^{i\frac{t}{n}X_i}]$$

$$= e^{-|t|}$$

This is the char. function of Cauchy.

Thus $\bar{X}_n \sim \text{Cauchy} \quad \forall n$

$$\Rightarrow \boxed{\bar{X}_n \xrightarrow{(d)} \text{Cauchy}}$$

Why don't we have $\bar{X}_n \xrightarrow{P} ?$

For this we have seen $\bar{X}_n \xrightarrow{P} \mu$ iff either

① $n P(|X_1| > n) \rightarrow 0$ + other conditions

Here
$$P(|X_1| > n) = \frac{2}{\pi} \int_n^\infty \frac{1}{1+x^2} dx$$
$$\geq c_1 \int_n^\infty \frac{1}{x^2} dx \geq c_2 \frac{1}{n}$$

Thus

$$n P(|X_1| > n) \not\rightarrow 0$$

② $\phi(t)$ is differentiable at 0
and $\phi'(0) = i\mu$

Here

$\phi(t) = e^{-|t|}$, not
differentiable at 0

Thus

$\bar{X}_n \xrightarrow{P} \star$

Next step Prove that

$$X_i \in L^1(\Omega) \Rightarrow \bar{X}_n \xrightarrow{a.s.} \mu = E[X_i]$$

Strategy: We already know

$$X_i \in L^2(\Omega) \Rightarrow \bar{X}_n \xrightarrow{a.s.} \mu$$

We will thus truncate the X_i 's
and then take limits

Truncation . Assume $X_j \geq 0$. Then
define

$$Y_n = X_n \mathbb{1}_{(X_n < n)}$$

Here truncation depends on n

Fact . Set

$$A_n = (X_n \neq Y_n)$$

Then

$$\begin{aligned} & P(A_n \text{ occurs i.o.}) \\ &= P(\limsup A_n) = 0 \end{aligned}$$

True if $\sum P(A_n) < \infty$
by B-C

Here

$$\begin{aligned} \sum_{n=1}^{\infty} P(A_n) &= \sum_{n=1}^{\infty} P(X_n \geq n) \\ &\stackrel{\text{i.i.d.}}{=} \sum_{n=1}^{\infty} P(X_1 \geq n) \\ &\stackrel{\text{p63}}{\leq} E[X_1] \stackrel{\text{L1}}{<} \infty \end{aligned}$$

Conclusion

$$P(X_n \neq Y_n \text{ i.o.}) = 0$$