

Discrete case

cond. pmf $p_{X|Y}(x|y) = \frac{p(x,y)}{p_Y(y)}$

Cond. exp $E[g(x) | Y=y] = \sum g(x) p_{X|Y}(x|y)$

Cont. case

Cond. density $f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)}$

Cond. exp $E[g(x) | Y=y]$
 $= \int_{\mathbb{R}} g(x) f_{X|Y}(x|y) dx$

Simple example of continuous conditioning (1)

Density: Let (X, Y) be a random vector with density

$$\frac{e^{-\frac{x}{y}} e^{-y}}{y} \mathbf{1}_{(0, \infty)}(x) \mathbf{1}_{(0, \infty)}(y) \quad (\Rightarrow \begin{matrix} x \geq 0 \\ y \geq 0 \text{ a.s.} \end{matrix})$$

Question: Compute

$$\begin{aligned} \mathbf{P}(X > 1 \mid Y = y) &= \mathbb{E}[\mathbf{1}_{(1, \infty)}(x) \mid Y = y] \\ &= \int_1^\infty f_{X|Y}(x|y) \, dx \end{aligned}$$

Marginal for Y : Take $y > 0$

$$\begin{aligned} f_Y(y) &= \int_{\mathbb{R}} f(x, y) \, dx \\ &= \int_{\mathbb{R}} \frac{e^{-y}}{y} e^{-\frac{x}{y}} \mathbb{1}_{(0, \infty)}(x) \, dx \\ &= e^{-y} \int_0^{\infty} \frac{1}{y} e^{-\frac{x}{y}} \, dx \\ &= e^{-y} \end{aligned}$$

Density of $\mathcal{E}(\frac{1}{y})$

($\Rightarrow Y \sim \mathcal{E}(1)$)

Cond. density Take $x, y > 0$

$$\begin{aligned} f_{X|Y}(x|y) &= \frac{f(x, y)}{f_Y(y)} = \frac{e^{-y} e^{-\frac{x}{y}}}{e^{-y} y} \\ &= \frac{1}{y} e^{-\frac{x}{y}} \end{aligned}$$

$$z \sim \mathcal{E}(1) \Rightarrow \mathbb{P}(z \geq a) = e^{-1a}$$

Summary For $x, y > 0$

$$f_{X|Y}(x|y) = \frac{1}{y} e^{-\frac{x}{y}}$$

$$\Rightarrow \mathcal{L}(X|Y=y) \sim \mathcal{E}\left(\frac{1}{y}\right)$$

Cond. probability

$$\begin{aligned} \mathbb{P}(X > 1 | Y = y) &= \mathbb{P}(z > 1) \quad \xrightarrow{\quad} z \sim \mathcal{E}\left(\frac{1}{y}\right) \\ &= e^{-\frac{1}{y}} \end{aligned}$$

Differences between baby / non baby cond. expectation

- (i) We condition on a σ -algebra $\mathcal{F}$
 $E[X | \mathcal{F}]$
- (ii) $E[X | \mathcal{F}]$ is a r.v.
- (iii) We define $E[g(x) | \mathcal{F}]$, then
conditional distributions
- (iv) Interpretation: "Best possible" approx.

Simple example of continuous conditioning (2)

Marginal distribution of Y : We have

$$\begin{aligned}f_Y(y) &= \int_0^\infty f(x, y) dx \\&= \frac{e^{-y}}{y} \left(\int_0^\infty e^{-\frac{x}{y}} dx \right) \mathbf{1}_{(0, \infty)}(y) \\&= e^{-y} \mathbf{1}_{(0, \infty)}(y)\end{aligned}$$

Conditional density: For $y > 0$ we have

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \frac{e^{-\frac{x}{y}}}{y} \mathbf{1}_{(0, \infty)}(x)$$

Namely $\mathcal{L}(X|Y = y) = \mathcal{E}(\frac{1}{y})$

Simple example of continuous conditioning (3)

Conditional probability:

$$\begin{aligned}\mathbf{P}(X > 1 | Y = y) &= \int_1^{\infty} f_{X|Y}(x|y) dx \\ &= \int_1^{\infty} \frac{e^{-\frac{x}{y}}}{y} dx \\ &= e^{-\frac{1}{y}}\end{aligned}$$

Cond. expectation in the continuous case

Definition 5.

Let

- (X, Y) couple of continuous random variables
- Joint density f
- Marginal densities f_X, f_Y , y such that $f_Y(y) > 0$
- $f_{X|Y}(x|y)$ conditional density

Then the conditional exp. of X given $Y = y$ is defined by

$$\mathbf{E}[X|Y = y] = \int_{\mathbb{R}} x f_{X|Y}(x|y) dx$$

Example of continuous conditional expectation (1)

Density: Let (X, Y) be a random vector with density

$$\frac{e^{-\frac{x}{y}} e^{-y}}{y} \mathbf{1}_{(0, \infty)}(x) \mathbf{1}_{(0, \infty)}(y)$$

Question: Compute

$$\mathbf{E}[X | Y = y]$$

Example of continuous conditional expectation (2)

Conditional density: For $y > 0$ we have seen that

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_Y(y)} = \frac{e^{-\frac{x}{y}}}{y} \mathbf{1}_{(0, \infty)}(x)$$

Namely $\mathcal{L}(X|Y = y) = \mathcal{E}(\frac{1}{y})$

Conditional expectation: We have

$$\begin{aligned} \mathbf{E}[X|Y = y] &= \int_{\mathbb{R}} x f_{X|Y}(x|y) dx \\ &= \int_0^{\infty} x \frac{e^{-\frac{x}{y}}}{y} \\ &= y \end{aligned}$$

Outline

1 Definition

- Baby conditional distributions: discrete case
- Baby conditional distributions: continuous case
- Definition with measure theory

2 Examples

3 Existence and uniqueness

4 Conditional expectation: properties

5 Conditional expectation as a projection

6 Conditional regular laws

Formal definition

Big σ -algebra

Definition 6.

We are given a probability space $(\Omega, \mathcal{F}_0, \mathbf{P})$ and

- A σ -algebra $\mathcal{F} \subset \mathcal{F}_0$.
- $X \in \mathcal{F}_0$ such that $\mathbf{E}[|X|] < \infty$.

Conditional expectation of X given $\mathcal{F}$:

- Denoted by $\mathbf{E}[X|\mathcal{F}]$
- Defined by: $\mathbf{E}[X|\mathcal{F}]$ is the $L^1(\Omega)$ r.v Y such that
 - (i) $Y \in \mathcal{F}$.
 - (ii) For all $A \in \mathcal{F}$, we have

$$\mathbf{E}[X|\mathcal{F}] = Y$$

$$\mathbf{E}[X1_A] = \mathbf{E}[Y1_A],$$

or otherwise stated $\int_A X d\mathbf{P} = \int_A Y d\mathbf{P}$.

Remarks

Notation: We use the notation $Y \in \mathcal{F}$ to say that a random variable Y is $\mathcal{F}$ -measurable.

Interpretation: More intuitively

- $\mathcal{F}$ represents a given information
- Y is the best prediction of X given the information in $\mathcal{F}$.

Existence and uniqueness:

To be seen after the examples.

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Setting $(\Omega, \mathcal{F}_0, P)$, $\mathcal{G} \subset \mathcal{F}_0$

Example 1 If $X \in \mathcal{G}$

Recall: Y is the "best approx"
of X as a r.v. $Y \in \mathcal{G}$

Intuitively: If $X \in \mathcal{G}$, the
best approx should be X

Hyp: $X \in \mathcal{F}$. Set $Y = X$. Then

(i) $Y \in \mathcal{F}$

(ii) If $A \in \mathcal{F}$

$$E[Y 1_A] = E[X 1_A]$$

$Y = X$

Thus

$$X = E[X | \mathcal{F}]$$

Easy examples (1)

Example 1: Assume

$$X \in \mathcal{F}.$$

Then

$$\mathbf{E}[X|\mathcal{F}] = X$$

Independence of a r.v and a σ -field

Definition 7.

We say that $X \perp\!\!\!\perp \mathcal{F}$ if

$\hookrightarrow$ for all $A \in \mathcal{F}$ and $B \in \mathcal{B}(\mathbb{R})$, we have

$$\mathbf{P}((X \in B) \cap A) = \mathbf{P}(X \in B) \mathbf{P}(A),$$

or otherwise stated:

$$X \perp\!\!\!\perp \mathbf{1}_A$$

Hyp: Consider $X \perp\!\!\!\perp F$. Then

$$E[X | F] = E[X]$$

Proof: Set $Y = E[X]$. Then

(i) $Y \in F$ (Y is constant)

(ii) If $A \in F$. Then

$$\begin{aligned} E[X \mathbb{1}_A] &\stackrel{||}{=} E[X] E[\mathbb{1}_A] \\ &= E[Y \mathbb{1}_A] \end{aligned}$$

Easy examples (2)

Example 2: Assume

$$X \perp\!\!\!\perp \mathcal{F}.$$

Then

$$\mathbf{E}[X|\mathcal{F}] = \mathbf{E}[X]$$

Proof: example 2

We have

- (i) $\mathbf{E}[X] \in \mathcal{F}$ since $\mathbf{E}[X]$ is a constant.
- (ii) If $A \in \mathcal{F}$,

$$\mathbf{E}[X \mathbf{1}_A] = \mathbf{E}[X] \mathbf{E}[\mathbf{1}_A] = \mathbf{E}\left[\mathbf{E}[X] \mathbf{1}_A\right].$$

Discrete conditional expectation

Example 3: We consider

- $\{\Omega_j; j \geq 1\}$ partition of Ω such that $\mathbf{P}(\Omega_j) > 0$ for all $j \geq 1$.
- $\mathcal{F} = \sigma(\Omega_j; j \geq 1)$.

Then

$$\mathbf{E}[X|\mathcal{F}] = \sum_{j \geq 1} \frac{\mathbf{E}[X \mathbf{1}_{\Omega_j}]}{\mathbf{P}(\Omega_j)} \mathbf{1}_{\Omega_j} \equiv Y. \quad (1)$$

Setting: Partition $\{\Omega_j; j \geq 1\}$

$$Y = \sum_{j \geq 1} \beta_j \mathbb{1}_{\Omega_j}, \quad \beta_j = \frac{E[X \mathbb{1}_{\Omega_j}]}{P(\Omega_j)}$$

$$(i) \quad \mathcal{F} = \sigma(\Omega_j; j \geq 1)$$

$$\Rightarrow \mathbb{1}_{\Omega_j} \in \mathcal{F} \quad (\text{as a r.v.})$$

$$\Rightarrow \sum_{j \geq 1} \alpha_j \mathbb{1}_{\Omega_j} \in \mathcal{F} \quad \forall (\alpha_j)_{j \geq 1}$$

In particular, $Y \in \mathcal{F}$

$$Y = \sum_{j \geq 1} \beta_j 1_{\Omega_j}, \quad \beta_j = \frac{E[Y 1_{\Omega_j}]}{P(\Omega_j)} \Rightarrow A = \bigcup_{j \in J} \Omega_j$$

$$(ii) \quad E[Y 1_A] \stackrel{(*)}{=} E[X 1_A] \quad \forall A \in \mathcal{F}$$

Note: Since $\{\Omega_j\}_{j \geq 1}$ is a partition, for every $A \in \mathcal{F}$ we have

$$1_A = \sum_{j \in J} 1_{\Omega_j} \quad J \subset \{1, 2, \dots\}$$

$\Rightarrow$ It is sufficient to prove $(*)$ for $A = \Omega_j, \quad j = 1, \dots$

$$\begin{aligned} \text{Then } E[Y 1_A] &= \sum_{j \in J} E[Y 1_{\Omega_j}] \\ &= \sum_{j \in J} E[X 1_{\Omega_j}] = E[X 1_A] \end{aligned}$$

$$Y = \sum_{j=1}^n \beta_j 1_{\Omega_j}, \quad \beta_j = \frac{E[X 1_{\Omega_j}]}{P(\Omega_j)}$$

For a given n ,

$= 0$ a.s. if $j \neq n$

$$E[X 1_{\Omega_n}] = E[(\sum_{j=1}^n \beta_j 1_{\Omega_j}) 1_{\Omega_n}]$$

$$= E[\beta_n 1_{\Omega_n}]$$

$$= \beta_n E[1_{\Omega_n}]$$

$$= \beta_n P(\Omega_n) = \frac{E[X 1_{\Omega_n}]}{P(\Omega_n)} P(\Omega_n)$$

$$= E[X 1_{\Omega_n}]$$

$\Rightarrow$ (ii) proved

Simple example of partition : B, B^c

$$\Rightarrow \sigma(B, B^c) = \{ \emptyset, B, B^c, \Omega \} = \mathcal{F}$$

Now take $A \in \mathcal{F}_0$ and set

$$E[1_A | \mathcal{F}] \equiv P(A | \mathcal{F})$$

We get, according to example 3

$$\begin{aligned} P(A | \mathcal{F}) &= \frac{E[1_A 1_B]}{P(B)} 1_B + \frac{E[1_A 1_{B^c}]}{P(B^c)} 1_{B^c} \\ &= P(A | B) 1_B + P(A | B^c) 1_{B^c} \end{aligned}$$

Proof: example 3

Strategy: Verify (i) and (ii) for the random variable Y .

(i) For all $j \geq 1$, we have $\mathbf{1}_{\Omega_j} \in \mathcal{F}$. Thus, for any sequence $(\alpha_j)_{j \geq 1}$,

$$\sum_{j \geq 1} \alpha_j \mathbf{1}_{\Omega_j} \in \mathcal{F}.$$

(ii) It is enough to verify (1) for $A = \Omega_n$ and $n \geq 1$ fixed. However,

$$\mathbf{E}[Y \mathbf{1}_{\Omega_n}] = \mathbf{E}\left\{ \frac{\mathbf{E}[X \mathbf{1}_{\Omega_n}]}{\mathbf{P}(\Omega_n)} \mathbf{1}_{\Omega_n} \right\} = \frac{\mathbf{E}[X \mathbf{1}_{\Omega_n}]}{\mathbf{P}(\Omega_n)} \mathbf{E}[\mathbf{1}_{\Omega_n}] = \mathbf{E}[X \mathbf{1}_{\Omega_n}].$$

Undergrad conditional probability

Definition: For a generic measurable set $A \in \mathcal{F}_0$, we set

$$\mathbf{P}(A|\mathcal{F}) \equiv \mathbf{E}[\mathbf{1}_A|\mathcal{F}]$$

Discrete example setting:

Let B, B^c be a partition of Ω , and $A \in \mathcal{F}_0$. Then

① $\mathcal{F} = \sigma(B) = \{\Omega, \emptyset, B, B^c\}$

② We have

$$\mathbf{P}(A|\mathcal{F}) = \mathbf{P}(A|B) \mathbf{1}_B + \mathbf{P}(A|B^c) \mathbf{1}_{B^c}.$$