

Outline

1 Introduction

- 1.1 Basic probability structures
- 1.2 Buffon's needle
- 1.3 Convergence of functions

2 Modes of convergence

- 2.1 Reviewing the modes of convergence
- 2.2 Results for P and L^p convergences
- 2.3 Results for almost sure convergence
- 2.4 Cases of inverse relations for modes of convergence
- 2.5 Inverse method for simulation
- 2.6 Results for convergence in distribution

Experiment



Procedure:

- Consider a plane ruled by lines $y = k$, with $k \in \mathbb{Z}$
- Take a needle with length 1
- Fling the needle n times on the plane

$$X_i = \mathbf{1}_{A_i} = \begin{cases} 1 & \text{if } A_i \text{ realized} \\ 0 & \text{otherwise} \end{cases}$$

Outcome: We record, for $i = 1, \dots, n$,

- $X_i \equiv \mathbf{1}_{A_i}$, where $A_i = (i\text{-th needle intersect a line})$
- $S_n \equiv \#$ times the needle intersects the line

Simulation:

This website from UIUC

Rmk We will see that

$\frac{S_n}{n}$ = average # needles intersecting

$$\xrightarrow{n \rightarrow \infty} \frac{2}{\pi}$$

This gives a way to approximate π

This can be seen as one of first
examples of Monte Carlo method
use of probability to compute
an integral

Limiting result

A_i = " i th needle hits one line "

Proposition 5.

Under the above conditions we have

$$\begin{aligned} \mathbf{P}(A_i) &= \frac{2}{\pi} \\ \frac{S_n}{n} &\longrightarrow \frac{2}{\pi} \end{aligned}$$

Bernoulli r.v

$$X \sim \mathcal{B}(p), \quad p \in (0,1)$$

$$X: \Omega \xrightarrow{\text{measurable}} \{0,1\}$$

Pmf:

$$p_X(0) \equiv \mathbb{P}(X=0) = 1-p$$

$$p_X(1) \equiv \mathbb{P}(X=1) = p$$

$$\mathbb{P}(\underbrace{\{\omega \in \Omega; X(\omega)=1\}}_{\in \mathcal{F}})$$

$\in \mathcal{F}$, since $\{1\}$ is measurable in \mathbb{R}

probability
mass
function

Bernoulli random variable (1)

Notation:

$$X \sim \mathcal{B}(p) \text{ with } p \in (0, 1)$$

State space:

$$\{0, 1\}$$

Pmf:

$$\mathbf{P}(X = 0) = 1 - p, \quad \mathbf{P}(X = 1) = p$$

Expected value, variance, generating function:

$$\mathbf{E}[X] = p, \quad \mathbf{Var}(X) = p(1 - p), \quad G_X(s) = (1 - p) + p s$$

Bernoulli random variable (2)

Bernoulli

Use 1, success in a binary game:

- Example 1: coin tossing
 - ▶ $X = 1$ if H, $X = 0$ if T
 - ▶ We get $X \sim \mathcal{B}(1/2)$
- Example 2: dice rolling
 - ▶ $X = 1$ if outcome = 3, $X = 0$ otherwise
 - ▶ We get $X \sim \mathcal{B}(1/6)$

17th

18th

19th

Use 2, answer yes/no in a poll

- $X = 1$ if a person feels optimistic about the future
- $X = 0$ otherwise
- We get $X \sim \mathcal{B}(p)$, with unknown p

Jacob Bernoulli

Some facts about Bernoulli:

- Lifespan: 1654-1705, in Switzerland
- Discovers constant e
- Establishes divergence of $\sum \frac{1}{n}$
- Contributions in diff. eq
- First law of large numbers
- Bernoulli:
family of 8 prominent mathematicians
- Fierce math fights between brothers



Binomial r.v

$$X \sim \text{Bin}(n, p)$$

$$X: \Omega \rightarrow \{0, 1, \dots, n\}$$

pmf: $k = 0, 1, \dots, n$

$$\mathbb{P}(X = k) = \binom{n}{k} p^k (1-p)^{n-k}$$

Binomial random variable (1)

Notation:

$$X \sim \text{Bin}(n, p), \text{ for } n \geq 1, p \in (0, 1)$$

State space:

$$\{0, 1, \dots, n\}$$

Pmf:

$$\mathbf{P}(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}, \quad 0 \leq k \leq n$$

Expected value, variance and generating function:

$$\mathbf{E}[X] = np, \quad \mathbf{Var}(X) = np(1 - p), \quad G_X(s) = [(1 - p) + ps]^n$$

Binomial random variable (2)

Use 1, Number of successes in a Bernoulli trial:

- Example: Roll a dice 9 times.
- $X = \#$ of 3 obtained
- We get $X \sim \text{Bin}(9, 1/6)$
- $\mathbf{P}(X = 2) = 0.28 = \binom{9}{2} \left(\frac{1}{6}\right)^2 \left(\frac{5}{6}\right)^7$

Use 2: Counting a feature in a repeated trial:

- Example: stock of 1000 pants with 10% defects
- Draw 15 times a pant at random
- $X = \#$ of pants with a defect
- We get $X \sim \text{Bin}(15, 1/10)$

Binomial random variable (3)

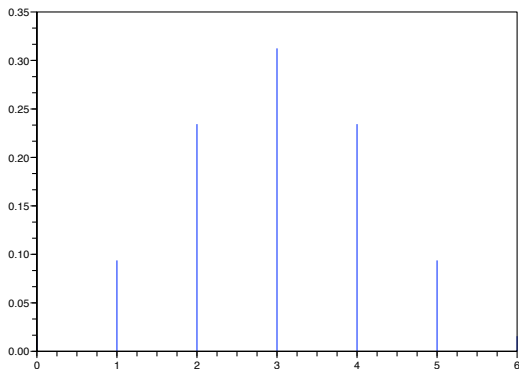


Figure: Pmf for $\text{Bin}(6; 0.5)$. x-axis: k . y-axis: $\mathbf{P}(X = k)$

Binomial random variable (4)

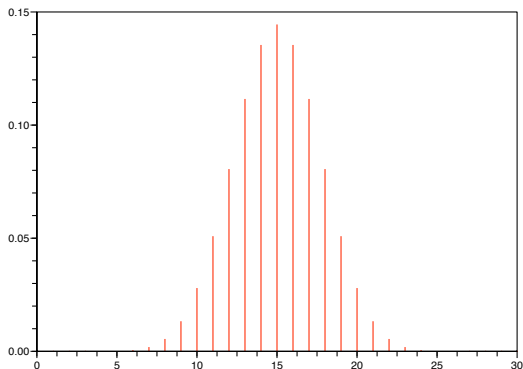


Figure: Pmf for $\text{Bin}(30; 0.5)$. x-axis: k . y-axis: $\mathbf{P}(X = k)$

Uniform random variable (1)

Notation:

$$X \sim \mathcal{U}([\alpha, \beta]), \text{ with } \alpha < \beta$$

State space:

$$[\alpha, \beta]$$

Density:

$$f(x) = \frac{1}{\beta - \alpha} \mathbf{1}_{[\alpha, \beta]}(x)$$

Expected value and variance:

$$\mathbf{E}[X] = \frac{\alpha + \beta}{2}, \quad \mathbf{Var}(X) = \frac{(\beta - \alpha)^2}{12}$$

Classical r.v

Discrete

$B(p)$

$\text{Bin}(n, p)$

$P(\lambda)$

$G(p)$

Continuous

$U([\alpha, \beta])$

$E(\lambda)$ \nearrow exponential
r.v

$N(\mu, \sigma^2)$

Cauchy

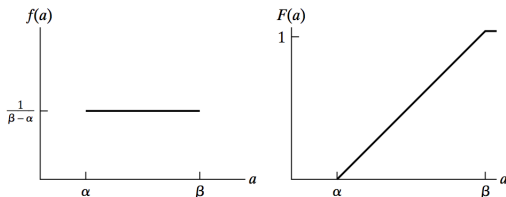
Uniform random variable (2)

Use:

- $\mathcal{U}([0, 1])$ only r.v directly accessible on a computer
 \hookrightarrow rand function

Example of computation: if $X \sim \mathcal{U}([8, 10])$, then

$$\mathbf{P}(7.5 < X < 9.5) = \frac{1}{2} \int_8^{9.5} dx = \frac{9.5 - 8}{2} = \frac{3}{4}$$



Experiment (repeated)

Procedure:

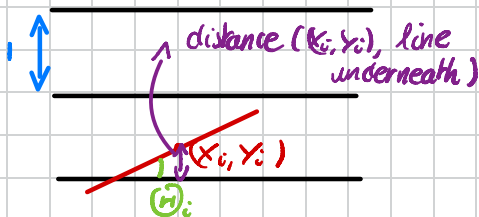
- Consider a plane ruled by lines $y = k$, with $k \in \mathbb{Z}$
- Take a needle with length 1
- Fling the needle n times on the plane

Outcome: We record, for $i = 1, \dots, n$,

- $X_i \equiv \mathbf{1}_{A_i}$, where $A_i = (i\text{-th needle intersects a line})$
- $S_n \equiv \#$ times the needle intersects the line

Model . Call

(X_i, Y_i) = center of
 i -th needle



θ_i = angle (i -th needle, x axis)

z_i = distance ((X_i, Y_i) , line underneath)

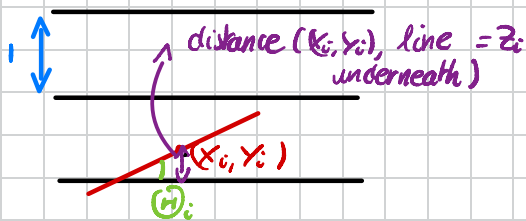
We assume independent and identically distributed

• $z_i \sim U([0, 1])$ • $z_i \perp \theta_i$

• $\theta_i \sim U([0, \pi])$ • $\{z_i; i \geq 1\}$ i.i.d
• $\{\theta_i; i \geq 1\}$ i.i.d

Recall:

A_i = "i-th needle hits a line"



We decompose

$$A_i = A_i^- \cup A_i^+ \quad \text{where}$$

$$A_i^- = \{ z_i \leq \frac{1}{2} \} \cap \{ z_i \leq \frac{1}{2} \sin \Theta_i \}$$

$$A_i^+ = \{ z_i > \frac{1}{2} \} \cap \{ 1 - z_i \leq \frac{1}{2} \sin \Theta_i \}$$

$$\Rightarrow \mathbb{P}(A_i) = \mathbb{P}(A_i^+) + \mathbb{P}(A_i^-) = 2 \mathbb{P}(A_i^-)$$

$$A_i^- = \{z_i \leq \frac{1}{2}\} \cap \{z_i \leq \frac{1}{2} \sin \Theta_i\}$$

$$z_i \sim \mathcal{U}([0,1]) \quad \Theta_i \sim \mathcal{U}([0,\pi]) \quad \underline{z_i \perp \Theta_i}$$

$$\mu(z_i, \Theta_i) = \mu_{z_i} \otimes \mu_{\Theta_i}$$

Thus

$$\begin{aligned} \boxed{\mathbb{P}(A_i^-)} &= \int_0^\pi d\theta \frac{1}{\pi} \int_0^1 dz \mathbb{1}(z \leq \frac{1}{2} \sin \theta) \\ &= \frac{1}{\pi} \int_0^\pi d\theta \int_0^{\frac{1}{2} \sin \theta} dz \\ &= \frac{1}{\pi} \int_0^\pi \frac{1}{2} \sin \theta d\theta = \frac{1}{2\pi} [-\cos \theta]_0^\pi \\ &= \frac{1}{\pi} \Rightarrow \mathbb{P}(A_i) = 2 \mathbb{P}(A_i^-) = \boxed{\frac{2}{\pi}} \end{aligned}$$

2nd step: Count the # hit. Set

$$P(A_i) = \frac{2}{\pi}$$

$$X_i = \mathbb{1}_{A_i} = \begin{cases} 1 & \text{if } i\text{-th hits a line} \\ 0 & \text{otherwise} \end{cases}$$

$$\Rightarrow X_i \sim B(p = \frac{2}{\pi})$$

$$S_n = \text{total number of hits} = \sum_{i=1}^n X_i$$

$$\Rightarrow S_n \sim \text{Bin}(n, \frac{2}{\pi})$$

According to de Moivre,

$$\boxed{\frac{S_n}{n} \rightarrow \frac{2}{\pi}}$$

Proof of Proposition 5 (1)

Notation: We define

- $(X_i, Y_i) \equiv$ Coordinates of the center of the i -th needle
- $\Theta_i \equiv$ angle (i -th needle, x-axis)
- $Z_i = d((X_i, Y_i), \text{nearest line underneath}) = Y_i - \lfloor Y_i \rfloor$

Model: We assume

- 1 $Z_i \sim \mathcal{U}([0, 1])$
- 2 $\Theta_i \sim \mathcal{U}([0, \pi])$
- 3 $Z_i \perp\!\!\!\perp \Theta_i$
- 4 $\{Z_i; i \geq 1\}$ i.i.d sequence
- 5 $\{\Theta_i; i \geq 1\}$ i.i.d sequence

Proof of Proposition 5 (2)

Expression for A_i : We have

$$A_i = A_i^- \cup A_i^+$$

with

$$A_i^- = \left\{ Z_i \leq \frac{1}{2}, \text{ and } Z_i < \frac{1}{2} \sin(\Theta_i) \right\}$$

$$A_i^+ = \left\{ Z_i > \frac{1}{2}, \text{ and } 1 - Z_i < \frac{1}{2} \sin(\Theta_i) \right\}$$

Proof of Proposition 5 (3)

Computing $\mathbf{P}(A_i)$: We write

$$\begin{aligned}\mathbf{P}(A_i) &= \mathbf{P}(A_i^-) + \mathbf{P}(A_i^+) \\ &= 2\mathbf{P}(A_i^-) \\ &= \frac{2}{\pi} \int_0^\pi d\theta \int_0^{\frac{1}{2}\sin(\theta)} dz\end{aligned}$$

Thus

$$\mathbf{P}(A_i) = \frac{2}{\pi}$$

Proof of Proposition 5 (4)

Some laws: We have

$$X_i \sim \mathcal{B}\left(\frac{2}{\pi}\right)$$

$$S_n \sim \text{Bin}\left(n, \frac{2}{\pi}\right)$$

Limit: By De Moivre,

$$\frac{S_n}{n} \longrightarrow \frac{2}{\pi}$$

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Aim of this chapter

Problem with limit statement:

- For every $n \geq 1$, we have $S_n : \Omega \rightarrow \mathbb{R}$
- S_n is thus a function
- We don't know exactly what $\frac{S_n}{n} \longrightarrow \frac{2}{\pi}$ means!

Aim of this chapter:

- Explore different modes of convergence for random variables

Preliminary step:

- Explore different modes of convergence for functions

Setting for convergence of functions

Sequence of functions: We consider

- A sequence $\{f_n; n \geq 1\}$ with

$$f_n : [0, 1] \longrightarrow \mathbb{R}$$

Aim of subsection: Review modes for

$$\lim_{n \rightarrow \infty} f_n$$

Pointwise convergence

Definition 6.

Let

- $\{f_n; n \geq 1\}$ sequence of measurable functions

We assume

$$u_n \longrightarrow \ell$$

$$\lim_{n \rightarrow \infty} f_n(x) = f(x), \text{ for all } x \in [0, 1]$$

Then we say that

$$f_n \longrightarrow f \text{ pointwise}$$

Almost everywhere convergence

Definition 7.

Let

- $\{f_n; n \geq 1\}$ sequence of measurable functions

We assume

$$(\Omega, \mathcal{F}, \mathbb{P})$$

$$\lim_{n \rightarrow \infty} f_n(x) = f(x), \text{ for almost every } x \in [0, 1]$$

Then we say that

$$f_n \longrightarrow f \text{ almost everywhere}$$

Question: what do we mean by
 $\lim_{n \rightarrow \infty} f_n(x) = f(x)$ for almost every $x \in [0,1]$?

Answer: If $\lambda \equiv$ Lebesgue measure on $[0,1]$,
 $\lambda(\{x \in [0,1]; f_n(x) \text{ does not converge to } f(x)\}) = 0$

L^p convergence

Definition 8.

Let

- $\{f_n; n \geq 1\}$ sequence of measurable functions

We assume

$$\|f_n - f\|_{L^p} = \left(\int |f_n(x) - f(x)|^p dx \right)^{1/p}$$

$$\lim_{n \rightarrow \infty} \|f_n(\cancel{x}) - f(\cancel{x})\|_{L^p([0,1])} = 0$$

Then we say that

$$f_n \longrightarrow f \text{ in } L^p([0,1])$$

Convergence in measure

Definition 9.

Let

- $\{f_n; n \geq 1\}$ sequence of measurable functions

We assume that for all $\varepsilon > 0$

$$\lim_{n \rightarrow \infty} \lambda(\{u \in [0, 1]; |f_n(u) - f(u)| > \varepsilon\}) = 0$$

Then we say that

$$f_n \longrightarrow f \text{ in measure}$$