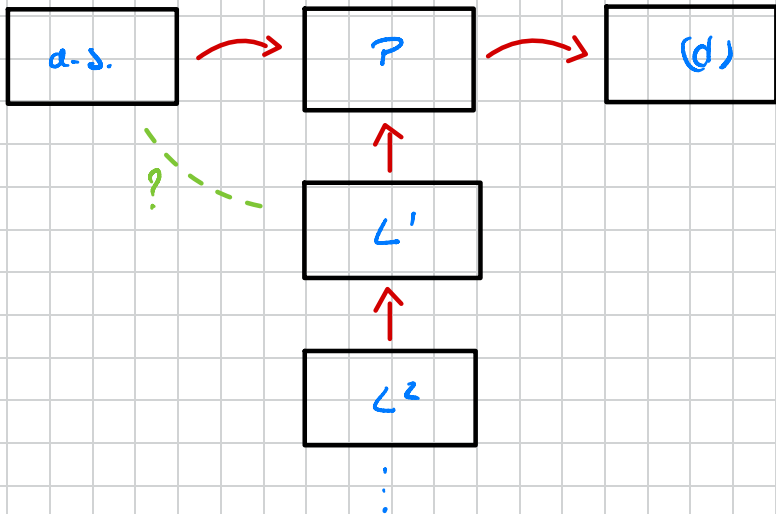


Summary



Non comparison between a.s and L^1 -convergence

Proposition 23.

One can find

- ① $\{X_n; n \geq 1\}$ sequence of random variables such that

$$X_n \xrightarrow{\text{a.s.}} X, \quad \text{but} \quad X_n \not\xrightarrow{L^1} X$$

- ② $\{Y_n; n \geq 1\}$ sequence of random variables such that

$$Y_n \xrightarrow{L^1} Y, \quad \text{but} \quad Y_n \not\xrightarrow{\text{a.s.}} Y$$

Counter-example 1 $(X_n)_{n \geq 1}$ of i.i.d. r.v. with

$$\mathbb{P}(X_n = n^3) = \frac{1}{n^2}, \quad \mathbb{P}(X_n = 0) = 1 - \frac{1}{n^2}$$

Then

(i) $X_n \xrightarrow{\text{a.s.}} 0$. To prove this, consider

$$A_n(\varepsilon) = \{ |X_n| > \varepsilon \}$$

It is sufficient to prove that $\forall \varepsilon > 0$,

$$\sum_{n=1}^{\infty} \mathbb{P}(A_n(\varepsilon)) < \infty$$

$$\mathbb{P}(X_n = n^3) = \frac{1}{n^2}, \quad \mathbb{P}(X_n = 0) = 1 - \frac{1}{n^2}$$

Here, with $x=0$, and $\varepsilon > 0$,

$$\mathbb{P}(A_n(\varepsilon)) = \mathbb{P}(|X - X_n| > \varepsilon)$$

$$= \mathbb{P}(X_n > \varepsilon) = \mathbb{P}(X_n = n^3) = \frac{1}{n^2}$$

Thus

$$\sum_{n \geq 1} \mathbb{P}(A_n(\varepsilon)) = \sum \frac{1}{n^2} < \infty$$

We get

$$X_n \xrightarrow{a.s.} 0$$

$$P(X_n = n^3) = \frac{1}{n^2}, \quad P(X_n = 0) = 1 - \frac{1}{n^2}$$

(ii) $X_n \xrightarrow{L^1} 0$. Indeed

$$E[|X_n - 0|] = E[X_n]$$

$$= 0 \times \left(1 - \frac{1}{n^2}\right) + n^3 \times \frac{1}{n^2}$$

$$= n \rightarrow 0$$

Counter example 2
sequence of i.i.d. r.v

Consider $(X_n)_{n \geq 1}$
with

$$X_n \sim \text{B}\left(\frac{1}{n}\right)$$

Then

(i) We have seen that $X_n \xrightarrow{\text{a.s.}} 0$

(ii) $X_n \xrightarrow{L^1} 0$, since

$$\mathbb{E}[|X_n - 0|] = \mathbb{E}[X_n] = \frac{1}{n} \xrightarrow{n \rightarrow \infty} 0$$

Proof of counter-example for X_n (1)

Definition of a sequence (repeated):

We consider independent r.v with

$$\mathbf{P}(X_n = n^3) = \frac{1}{n^2}, \quad \mathbf{P}(X_n = 0) = 1 - \frac{1}{n^2}$$

Convergence: We have

- 1 $X_n \xrightarrow{\text{a.s.}} 0$
- 2 X_n does not converge in L^1

Proof of counter-example for X_n (2)

Some notation: For $\varepsilon > 0$ set:

$$A_k(\varepsilon) = \{|X_k - X| > \varepsilon\}$$

Almost sure convergence: We have

$$\begin{aligned}\sum_{n=1}^{\infty} \mathbf{P}(A_n(\varepsilon)) &= \sum_{n=1}^{\infty} \mathbf{P}(X_n = n^3) \\ &= \sum_{n=1}^{\infty} \frac{1}{n^2} \\ &< \infty\end{aligned}$$

Thus

$$X_n \xrightarrow{\text{a.s.}} 0$$

Proof of counter-example for X_n (3)

Non convergence in $L^1(\Omega)$: We have already seen that

$$\mathbf{E}[|X_n|] = \mathbf{E}[X_n] = n$$

Thus

$$X_n \not\stackrel{L^1}{\rightarrow} 0$$

Outline

1 Introduction

- 1.1 Basic probability structures
- 1.2 Buffon's needle
- 1.3 Convergence of functions

2 Modes of convergence

- 2.1 Reviewing the modes of convergence
- 2.2 Results for P and L^p convergences
- 2.3 Results for almost sure convergence
- 2.4 Cases of inverse relations for modes of convergence
- 2.5 Inverse method for simulation
- 2.6 Results for convergence in distribution

Case for which $\xrightarrow{(d)}$ yields $\xrightarrow{P}$

Proposition 24.

Consider

- $\{X_n; n \geq 1\}$ sequence of random variables

Assume

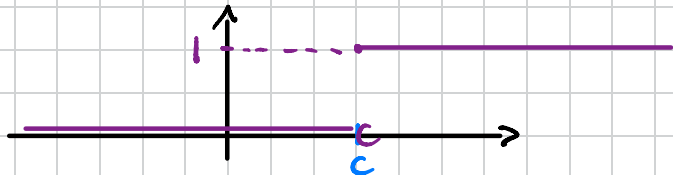
$$X_n \xrightarrow{(d)} c, \text{ where } c \text{ is a constant}$$

Then we have:

$$X_n \xrightarrow{P} c$$

Hyp $\overbrace{P(X_n \leq x)}^{F_n(x)} \rightarrow \overbrace{P(X \leq x)}^{F(x)}$ for every point of continuity of F

Hyp 2 $X = c \Rightarrow F(x) = 1_{[c, \infty)}(x)$



Aim: Prove $P(|X_n - c| > \varepsilon) \xrightarrow{n \rightarrow \infty} 0$
 $\forall \varepsilon > 0$

Computation



$$P(|X - X_n| > \epsilon) = P(|X_n - c| > \epsilon)$$

$$= P(X_n < c - \epsilon) + P(X_n > c + \epsilon)$$

$$= P(X_n < c - \epsilon) + 1 - P(X_n \leq c + \epsilon)$$

$$\leq P(X_n \leq c - \epsilon)$$

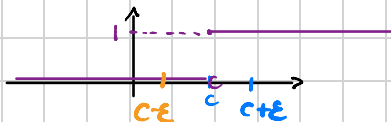
$\downarrow$
 $\downarrow n \rightarrow \infty$

$$F(c - \epsilon) = 0$$

$$F(c + \epsilon) \xrightarrow{n \rightarrow \infty} 1$$

Conclusion: As $n \rightarrow \infty$,
 $P(|X - X_n| > \epsilon) \rightarrow 0$

$$\Rightarrow \boxed{X_n \xrightarrow{P} c}$$



Proof of Proposition 24

Expression in terms of cdf: We have

$$\begin{aligned}\mathbf{P}(|X_n - c| > \varepsilon) &= \mathbf{P}(X_n < c - \varepsilon) + \mathbf{P}(X_n > c + \varepsilon) \\ &= \mathbf{P}(X_n < c - \varepsilon) + 1 - \mathbf{P}(X_n \leq c + \varepsilon)\end{aligned}$$

Convergence: Since $X_n \xrightarrow{(d)} X$, we get

$$\lim_{n \rightarrow \infty} \mathbf{P}(|X_n - c| > \varepsilon) = 0$$

Case for which $\xrightarrow{P}$ yields $\xrightarrow{\text{a.s.}}$

Proposition 25.

Consider

- $\{X_n; n \geq 1\}$ sequence of random variables

Assume

$$X_n \xrightarrow{P} X$$

Then there exists a subsequence $\{n_k; k \geq 1\}$ such that:

$$X_{n_k} \xrightarrow{\text{a.s.}} X$$

Proof Consider event $(|X - x_n| > \varepsilon)$
we know $P(|X - x_n| > \varepsilon) \rightarrow 0 \quad \forall \varepsilon$

Define a subsequence recursively: $n_1 = 1$
and

$$n_k = \inf \left\{ n > n_{k-1} ; P(|X_n - X| > \frac{1}{k}) \leq \frac{1}{k^2} \right\}$$

Subsequence :

$$Y_k = X_{n_k}$$

Consider $A_k(\varepsilon) = \{|Y_k - X| > \varepsilon\}$

Aim: prove that $\sum_{k \geq \frac{1}{\varepsilon}} P(A_k(\varepsilon)) < \infty$

Recall: $P(|Y_k - x| > \frac{1}{k}) \leq \frac{1}{k^2}$

Computation

$$\sum_{k \geq \frac{1}{\varepsilon}} P(A_k(\varepsilon)) \leq \sum_{k \geq \frac{1}{\varepsilon}} \frac{1}{k^2} < \infty$$

Conclusion

$$Y_k \xrightarrow{\text{a.s.}} x$$

Proof of Proposition 25 (1)

Definition of n_k : Recursively we set

$$n_k = \inf \left\{ n > n_{k-1}; \mathbf{P} \left(|X_n - X| > \frac{1}{k} \right) \leq \frac{1}{k^2} \right\}$$

Some notation: For $\varepsilon > 0$ define:

$$\begin{aligned} Y_k &= X_{n_k} \\ A_k(\varepsilon) &= \{|Y_k - X| > \varepsilon\} \end{aligned}$$

Proof of Proposition 25 (2)

Almost sure convergence: We have

$$\begin{aligned}\sum_{k=\varepsilon^{-1}}^{\infty} \mathbf{P}(A_n(\varepsilon)) &= \sum_{k=\varepsilon^{-1}}^{\infty} \mathbf{P}\left(|X_{n_k} - X| > \frac{1}{k}\right) \\ &\leq \sum_{k=1}^{\infty} \frac{1}{k^2} \\ &< \infty\end{aligned}$$

Thus

$$Y_k \xrightarrow{\text{a.s.}} X$$

Case for which $\xrightarrow{P}$ yields $\xrightarrow{L^r}$

Proposition 26.

Consider

- $\{X_n; n \geq 1\}$ sequence of random variables

Assume (bounded convergence) $\xrightarrow{\text{better?}} \text{dominated convergence}$

$X_n \xrightarrow{P} X$, and $|X_n| \leq k$ a.s for all $n \geq 1$ and a given $k > 0$

Then for all $r \geq 1$ we have:

$$X_n \xrightarrow{L^r} X$$

Proof - 1st step : We know that a.s.,
 $|X_n| \leq k \quad \forall n$. We can prove that
 X is also bounded. To this aim:

$$B_\delta = (|X| \leq k + \delta) \Rightarrow \mathbb{P}(B_\delta) = 1$$

Then

$$\begin{aligned} \mathbb{P}(B_\delta^c) &= \mathbb{P}(|X| > k + \delta) \quad \mathbb{P}(|X_n| \leq k) = 1 \\ &= \mathbb{P}(|X| > k + \delta, |X_n| \leq k) \\ &\leq \mathbb{P}(|X - X_n| > \delta) \xrightarrow{n \rightarrow \infty} 0 \quad \forall \delta > 0 \end{aligned}$$

Taking $\delta \rightarrow 0$, we get $\mathbb{P}(|X| \leq k) = 1$

Step 2: Assume $X_n \xrightarrow{P} X$, $|X_n| \leq k$. Then

$$\begin{aligned} E[|X - X_n|^r] &= E[|X - X_n|^r \mathbb{1}_{(|X - X_n| > \varepsilon)}] \\ &\quad + E[|X - X_n|^r \mathbb{1}_{(|X - X_n| \leq \varepsilon)}] \end{aligned}$$

$$\leq (2k)^r P(|X - X_n| > \varepsilon) + \varepsilon^r$$

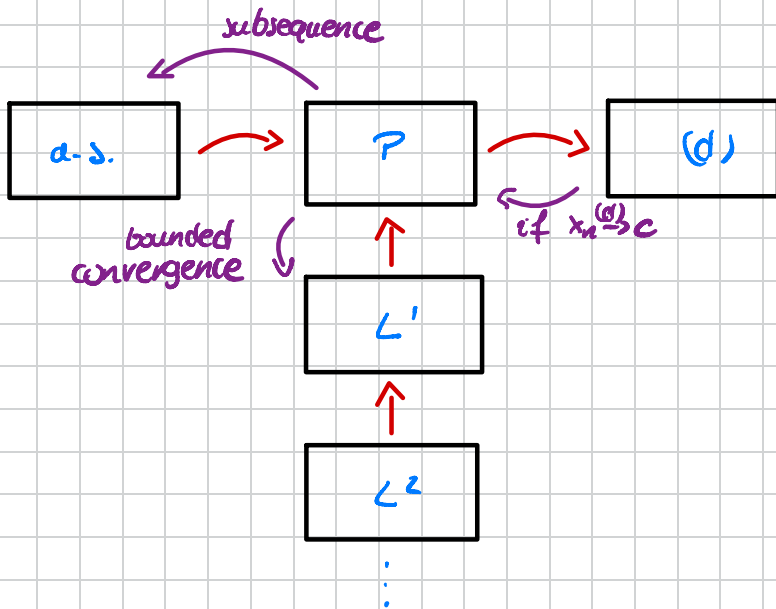
$\rightarrow \varepsilon^r$ as $n \rightarrow \infty$, for all $\varepsilon > 0$

Conclusion

Def: $E[|X - X_n|^r] \xrightarrow{n \rightarrow \infty} 0$

$$X_n \xrightarrow{L^r} X$$

New picture



Proof of Proposition 26 (1)

Boundedness of X : For $\delta > 0$, set

$$B_\delta = (|X| \leq k + \delta)$$

Then for all $n \geq 1$ we have

$$\begin{aligned}\mathbf{P}(B_\delta) &\geq \mathbf{P}(|X - X_n| \leq \delta, |X_n| \leq k) \\ &\geq \mathbf{P}(|X_n| \leq k) - \mathbf{P}(|X - X_n| > \delta) \\ &= 1 - \mathbf{P}(|X - X_n| > \delta)\end{aligned}$$

Taking limits in n, δ we get

$$\mathbf{P}(|X| \leq k) = 1$$

Proof of Proposition 26 (2)

Decomposition of $X_n - X$: For $\varepsilon > 0$ and $n \geq 1$ set

$$A_{n,\varepsilon} = \{|X_n - X| > \varepsilon\}$$

Then

$$|X_n - X|^r \leq \varepsilon^r \mathbf{1}_{A_{n,\varepsilon}^c} + (2k)^r \mathbf{1}_{A_{n,\varepsilon}}$$

Taking expectations: We obtain

$$\mathbf{E}[|X_n - X|^r] \leq \varepsilon^r \mathbf{P}(A_{n,\varepsilon}^c) + (2k)^r \mathbf{P}(A_{n,\varepsilon})$$

Taking limits: With $n \rightarrow \infty$ and $\varepsilon \rightarrow 0$ we end up with

$$\lim_{n \rightarrow \infty} \mathbf{E}[|X_n - X|^r] = 0$$

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Right inverse (1)

Definition 27.

Let $F : \mathbb{R} \rightarrow [0, 1]$ continuous cdf

We define the **right inverse** F^{-1} as

$$F^{-1} : (0, 1) \rightarrow \mathbb{R}, \quad y \mapsto \inf \{a \in \mathbb{R}; F(a) \geq y\}$$

Right inverse (2)

Remarks on right inverse:

(i) If F is strictly increasing, F^{-1} is the inverse of F
 $\hookrightarrow$ i.e. $F \circ F^{-1} = F^{-1} \circ F = \text{Id}$

(ii) Graphical method to construct F^{-1} :

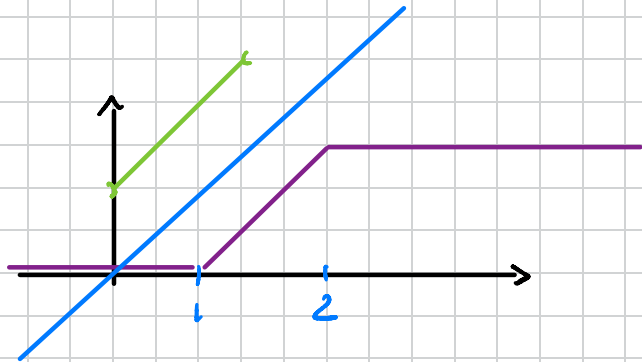
- 1 Symmetry wrt diagonal
- 2 Then erase vertical parts

Example: $F(x) = (x - 1)\mathbf{1}_{[1,2)}(x) + \mathbf{1}_{[2,\infty)}(x)$
 $\hookrightarrow F^{-1}(y) = (1 + y)\mathbf{1}_{(0,1)}(y)$

$$F \leftarrow U([1,2])$$

Example of right inverse

$$F(x) = (x-1) \mathbb{1}_{[1,2)}(x) + \mathbb{1}_{[2,\infty)}(x)$$



$$F^{-1}(y) = (y+1) \mathbb{1}_{(0,1)}(y)$$

Rmk F is the cdf of $U([1,2])$

$\Rightarrow [1,2]$ is what is relevant for the r.v.

On $(1,2)$, F is strictly monotone
 $\Rightarrow$ one can compute the inverse

$$F(x) = y \Leftrightarrow x - 1 = y$$

$$\Leftrightarrow x = y + 1 = F^{-1}(y)$$