Ergodic Theorems

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Probability Theory 2 - MA 539

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Outline

Definitions and examples

- Preliminaries on Markov chains
- Examples of stationary sequences
- Notion of ergodicity

2 Ergodic theorem



Overview

Stationary sequence: such that

$$\{X_{n+k}; n \ge 0\} \stackrel{(d)}{=} \{X_n; n \ge 0\}.$$

Main result: law of large numbers of the type

$$\frac{1}{n}\sum_{k=1}^{n}f(X_k) \quad \text{exists a.s}$$

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2 Ergodic theorem

3 Recurrence

< 1 k

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3 Recurrence

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Markov chain

Definition 1.

Let $X = \{X_n; n \ge 0\}$ be a process.

X is a Markov chain if

 $\mathbf{P}(X_{n+1} = j \mid X_0 = i_0, \dots, X_n = i_n) = \mathbf{P}(X_{n+1} = j \mid X_n = i_n)$

for all $n \geq 0$, $i_0, \ldots, i_n, j \in E$.

In the Markov chain is homogeneous whenever

$$\mathbf{P}(X_{n+1} = j | X_n = i) = \mathbf{P}(X_1 = j | X_0 = i)$$

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Law of the Markov chain

Proposition 2.

Let X be a homogeneous Markov chain \hookrightarrow with initial law ν and transition p.

• For $n \in \mathbb{N}$ and $i_0, \ldots, i_n \in E$, we have

 $\mathbf{P}(X_0 = i_0, \dots, X_n = i_n) = \nu(i_0) p(i_0, i_1) \cdots p(i_{n-1}, i_n)$

2 The law of X is characterized by ν and p.

Law of the Markov chain: general state space



A criterion for Markovianity

Proposition 4. l et • $Z: \Omega \rightarrow E$ random variable. • F countable set. • $\{Y_n; n \ge 1\}$ i.i.d sequence, with $Y \perp Z$ and $Y_n \in F$. • $f: E \times F \to E$. We set $X_0 = Z$, and $X_{n+1} = f(X_n, Y_{n+1})$. Then X is a homogeneous Markov chain such that $\nu_0 = \mathcal{L}(Z)$, and $p(i,j) = \mathbf{P}(f(i, Y_1) = j)$.

Remark: A converse result exists, but it won't be used.

Graph of a Markov chain

Definition 5.

Let X homogeneous Markov chain (initial law ν_0 , transition p). We define a graph $\mathcal{G}(X)$ by

- $\mathcal{G}(X)$ is an oriented graph
- Vertices of $\mathcal{G}(X)$ are points of E.
- Edges of $\mathcal{G}(X)$ are defined by

 $\mathbb{V} \equiv \{(i,j); i \neq j, p(i,j) > 0\}.$

Example

Example 6.

We consider $E = \{1, 2, 3, 4, 5\}$ and

$$p = \begin{pmatrix} 1/3 & 0 & 2/3 & 0 & 0 \\ 1/4 & 1/2 & 1/4 & 0 & 0 \\ 1/2 & 0 & 1/2 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 2/3 & 1/3 \end{pmatrix}$$

Related graph: to be done in class

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Graph and accessibility

Proposition 7.

Let X homogeneous Markov chain (initial law ν_0 , transition p). Then

$i \rightarrow j$ iff

i = j or there exists an oriented path from i to j in $\mathcal{G}(X)$.

Minimal class

Definition 8.

An equivalence class C is minimal if:

For all $i \in C$ and $j \notin C$, we have $i \not\rightarrow j$.

Example:

For Example 6, C_1 , C_3 are minimal, and C_2 is not minimal.

Minimality criterions:

(i) If there exists a unique class C, it is minimal.

(ii) There exists a unique minimal class C

 $\Leftrightarrow \exists ! \text{ class } C \text{ such that for all } i \in E, \text{ we have } i \to C.$

Application: Random walk

Recurrence criterions



Example

Recall: In Example 6 we had $E = \{1, 2, 3, 4, 5\}$ (hence $|E| < \infty$) and

$$p = \begin{pmatrix} 1/3 & 0 & 2/3 & 0 & 0 \\ 1/4 & 1/2 & 1/4 & 0 & 0 \\ 1/2 & 0 & 1/2 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 \\ 0 & 0 & 0 & 2/3 & 1/3 \end{pmatrix}$$

Related classes: $C_1 = \{1,3\}, C_2 = \{2\} \text{ and } C_3 = \{4,5\}.$ $\hookrightarrow C_1, C_3 \text{ minimal and } C_2 \text{ not minimal.}$

Conclusion: C_1 , C_3 positive recurrent, C_2 transient.

Outline

1 Definitions and examples

Preliminaries on Markov chains

• Examples of stationary sequences

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3 Recurrence

< 47 ▶

Stationary sequence

Definition 10.

Let

•
$$\{X_n; n \ge 0\}$$
 sequence of random variables

We say that X is stationary if for all $k \ge 1$ we have

$$\{X_{n+k}; n \ge 0\} \stackrel{(d)}{=} \{X_n; n \ge 0\}.$$

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iid random variables

Example 11.

Let

• $X = \{X_n; n \ge 0\}$ sequence of i.i.d random variables

Then X is a stationary sequence.

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Markov chains

Example 12. Let • $X = \{X_n; n \ge 0\}$ Markov chain on state space S• Transition probability: p(x, A)• Hypothesis 1: unique stationary distribution π

• Hypothesis 2:
$$\mathcal{L}(X_0) = \pi$$

Then X is a stationary sequence.

Proof:

Easy consequence of relation (1).

Trivial example of Markov chain

Example 13.
On
$$S = \{0, 1\}$$
 we take
• $p(x, 1 - x) = 1$
• $\pi(0) = \frac{1}{2}, \pi(1) = \frac{1}{2}$
Then
 $P(X = (0, 1, 0, 1, ...)) = P(X = (1, 0, 1, 0, ...)) = \frac{1}{2}$

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Rotation of the circle

Example 14.

We consider $\lambda \equiv$ Lebesgue measure and:

•
$$(\Omega, \mathcal{F}, \mathbf{P}) = ([0, 1], \text{ Borel sets}, \lambda)$$

•
$$heta \in (0,1)$$

•
$$X = \{X_n; n \ge 0\}$$
 with

 $X_n(\omega) = (\omega + n\theta) \mod 1 = \omega + n\theta - [\omega + n\theta].$

Then X is a stationary sequence.

Proof

Law of X_1 : For $\varphi \in \mathcal{C}_b(\mathbb{R})$ we have

$$\begin{split} \mathbf{\mathsf{E}}\left[\varphi(X_{1})\right] &= \int_{0}^{1} \varphi\left(x+\theta-\left[x+\theta\right]\right) \, dx \\ &= \int_{0}^{1-\theta} \varphi\left(x+\theta\right) \, dx + \int_{1-\theta}^{1} \varphi\left(x+\theta-1\right) \, dx \\ &= \int_{0}^{1} \varphi\left(v\right) \, dv = \mathbf{\mathsf{E}}\left[\varphi(X_{0})\right] \end{split}$$

Thus $\mathcal{L}(X_1) = \mathcal{L}(X_0)$.

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Proof (2)

Altenative expression for X_n : Due to the relation

 $(a+b) \mod 1 = [a \mod 1 + b \mod 1] \mod 1$,

we have

$$\begin{array}{rcl} X_n &=& (\omega + n\theta) \mod 1 \\ &=& (\omega + (n-1)\theta + \theta) \mod 1 \\ &=& [(\omega + (n-1)\theta) \mod 1 + \theta \mod 1] \mod 1 \\ &=& [X_{n-1} + \theta] \mod 1 \end{array}$$

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Proof (3)

Recall: We have seen that

$$X_n = [X_{n-1} + heta] \mod 1.$$

Conclusion: It is readily checked that

- From (2) and Proposition 4, X is a Markov chain
- **②** The relation $\mathcal{L}(X_1) = \mathcal{L}(X_0)$ means that λ is an invariant distribution

Therefore X is stationary thanks to Example 12.

(2)

Transformation of a stationary sequence

Theorem 15.

Let

- $X = \{X_n; n \ge 0\}$ stationary sequence
- $g: \mathbb{R}^{\mathbb{N}} \to \mathbb{R}$ measurable

We consider a sequence Y defined by

 $Y_k = g(\{X_{k+n}; n \ge 0\}).$

Then Y is a stationary sequence.

$$\sigma$$
-algebra on $\mathbb{R}^{\mathbb{N}}$ (1)

Set $\mathbb{R}^{\mathbb{N}}$: We define

$$\mathbb{R}^{\mathbb{N}} = \{ \omega = (\omega_j)_{j \in \mathbb{N}}; \, \omega_j \in \mathbb{R} \} \,.$$

Finite dimensional set: Of the form

$$A = \left\{ \omega \in \mathbb{R}^{\mathbb{N}}; \ \omega_j \in B_j, \ ext{for} \ 1 \leq j \leq n
ight\}, \quad ext{with} \quad B_j \in \mathcal{R},$$

where $\mathcal{R} \equiv$ Borel σ -algebra in \mathbb{R} .

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 σ -algebra on $\mathbb{R}^{\mathbb{N}}$ (2)

Definition 16.

On $\mathbb{R}^{\mathbb{N}}$ we set:

 $\mathcal{R}^{\mathbb{N}} = \sigma$ (finite dimensional sets).

Then $\mathcal{R}^{\mathbb{N}}$ is called Borel σ -algebra on $\mathbb{R}^{\mathbb{N}}$.

Remark:

Kolmogorov's extension theorem is valid on $(\mathbb{R}^{\mathbb{N}}, \mathcal{R}^{\mathbb{N}})$.

Proof of Theorem 15

Notation: For $x \in \mathbb{R}^{\mathbb{N}}$ and $k \ge 0$ we set

$$g_k(x) = g(\{x_{k+n}; n \ge 0\}).$$

Inverse image: Let $B \in \mathcal{R}^{\mathbb{N}}$. Then

$$\left\{\omega \in \mathbb{R}^{\mathbb{N}}; Y \in B\right\} = \left\{\omega \in \mathbb{R}^{\mathbb{N}}; X \in A\right\},$$

where

$$A = \left\{ \omega \in \mathbb{R}^{\mathbb{N}}; g(X) \in B \right\}.$$

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Image: A matrix

Proof of Theorem 15 (2)

Proof of stationarity: For the generic $B \in \mathcal{R}^{\mathbb{N}}$ we have

$$\begin{aligned} \mathbf{P}\left(\omega\in\mathbb{R}^{\mathbb{N}};\ Y\in B\right) &= &\mathbf{P}\left(\omega\in\mathbb{R}^{\mathbb{N}};\ X\in A\right) \\ &= &\mathbf{P}\left(\omega\in\mathbb{R}^{\mathbb{N}};\ \{X_{n};\ n\geq 0\}\in A\right) \\ &= &\mathbf{P}\left(\omega\in\mathbb{R}^{\mathbb{N}};\ \{X_{k+n};\ n\geq 0\}\in A\right) \\ &= &\mathbf{P}\left(\omega\in\mathbb{R}^{\mathbb{N}};\ \{Y_{k+n};\ n\geq 0\}\in B\right). \end{aligned}$$

This yields stationarity for Y.

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Image: A matrix

Bernoulli shifts

Example 17.

We consider $\lambda \equiv$ Lebesgue measure and:

•
$$(\Omega, \mathcal{F}, \mathbf{P}) = ([0, 1], \text{ Borel sets}, \lambda)$$

•
$$Y_0 = \omega$$

•
$$Y = \{Y_n; n \ge 0\}$$
 where for $n \ge 1$ we have

$$Y_n(\omega) = 2Y_{n-1} \mod 1.$$

Then Y is a stationary sequence.

Proof by Markov chains

Law of Y_1 : For $\varphi \in \mathcal{C}_b(\mathbb{R})$ we have

$$\mathbf{E} \left[\varphi(Y_1) \right] = \int_0^1 \varphi \left(2y \mod 1 \right) \, dy$$
$$= \int_0^{\frac{1}{2}} \varphi \left(2y \right) \, dy + \int_{\frac{1}{2}}^1 \varphi \left(2y - 1 \right) \, dy$$
$$= \int_0^1 \varphi \left(v \right) \, dv = \mathbf{E} \left[\varphi(Y_0) \right]$$

Thus $\mathcal{L}(Y_1) = \mathcal{L}(Y_0)$.

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Proof by Markov chains (2)

Proof of stationarity: We have

- Thanks to proposition 4, Y is thus a Markov chain
- The relation $\mathcal{L}(Y_1) = \mathcal{L}(Y_0)$ means that λ is an invariant distribution for Y

Therefore Y is a stationary sequence thanks to Example 12.

Proof by Theorem 15

Another representation for Y: Let

• {
$$X_i$$
; $i \ge 1$ } i.i.d with common law $\mathcal{B}(1/2)$
• $g(x) = \sum_{i\ge 1} x_i 2^{-(i+1)}$ defined for $x \in \{0,1\}^{\mathbb{N}}$
• $g_k(x) = g(\{x_{k+i}; i \ge 1\})$ defined for $x \in \{0,1\}^{\mathbb{N}}$
• $Y_k = g_k(X)$

Then

$$Y_0 \sim \lambda$$
, and $Y_n(\omega) = 2Y_{n-1} \mod 1$.

Stationarity of Y: We have

• X stationary

•
$$Y_k = g_k(X)$$

Therefore Y is a stationary sequence.

Measure preserving map



Measure preserving map and stationarity

Example 19.

Let

- $(\Omega, \mathcal{F}, \mathbf{P})$ a probability space
- A measure preserving map $\varphi: \Omega \to \Omega$
- $X \in \mathcal{F}$ a random variable

For $n \ge 0$ we set:

$$X_n = X\left(\varphi^n(\omega)\right)$$

Then X is a stationary sequence.

Proof

Characterization with expected values: For $g \in C(\mathbb{R}^n)$ and $k \ge 1$ we have:

$$\mathbf{E} [g (X_k, \dots, X_{k+n})] = \mathbf{E} \left[g \left(X_0(\varphi^k(\omega)), \dots, X_n(\varphi^k(\omega)) \right) \right]$$

=
$$\mathbf{E} [g (X_0(\omega), \dots, X_n(\omega))]$$

=
$$\mathbf{E} [g (X_0, \dots, X_n)].$$

Thus X is a stationary sequence.

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Image: A matrix
Stationarity and measure preserving maps

Example 20.

Let

- Y stationary sequence with values in $S = \mathbb{R}^n$
- **P** probability measure on $(S^{\mathbb{N}}, \mathcal{R}(S^{\mathbb{N}}))$ defined by:

$$X_n(\omega) = \omega_n \implies \mathcal{L}(X) = \mathcal{L}(Y)$$

We define a shift operator φ by:

$$\varphi\left(\{\omega_j; j \ge 0\}\right) = \{\omega_{j+1}; j \ge 0\}.$$

Then φ is measure preserving.

Interpretation:

Previous examples can be reduced to a measure preserving map

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Two sided stationary sequence

Theorem 21. Let • X stationary sequence Then X can be embedded into a two sided sequence $\{Y_n; n \in \mathbb{Z}\}.$

Proof

Application of Kolmogorov's extension:

It is enough to define a family of probability measures

- On \mathbb{R}^A for any finite subset of \mathbb{Z}
- With consistent property

Current situation: we consider Y defined by

$$E[g(Y_{-m},...,Y_n)] = E[g(X_0,...,X_{m+n})],$$

for all $g \in \mathcal{C}_b(\mathbb{R}^{m+n+1})$. Kolmogorov's extension applies

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3 Recurrence

< 47 ▶

Setup

General setting: We are reduced to

- $(\Omega, \mathcal{F}, \mathbf{P})$ probability space
- A map φ preserving ${\bf P}$
- A random variable X
- A sequence $X_n(\omega) = X(\omega)$

Invariant set: Let $A \in \mathcal{F}$. The set A is invariant if

 $\mathbf{P}\left(A\,\Delta\,\varphi^{-1}(A)
ight)=0$

Notation abuse: For an invariant set A, we often write

$$A = \varphi^{-1}(A)$$

Ergodicity

Definition 22.

Let

- $(\Omega, \mathcal{F}, \mathbf{P})$ probability space
- A map φ preserving **P**
- $\mathcal{I} \equiv \sigma$ -algebra of invariant events

We say that φ is ergodic if ${\mathcal I}$ is trivial, i.e.

 $A \in \mathcal{I} \implies \mathbf{P}(A) \in \{0,1\}.$

Kolmogorov's 0-1 law

Theorem 23.

Let

X = {X_n; n ≥ 0} sequence of i.i.d random variables
For n ≥ 0, we set F'_n = σ({X_k; k ≥ n})
Tail σ-field: T ≡ ∩_{n≥0}F'_n

Then \mathcal{T} is trivial, i.e.

 $A \in \mathcal{T} \implies \mathbf{P}(A) \in \{0,1\}.$

Ergodicity for i.i.d sequences

Example 24.
Let
•
$$X = \{X_n; n \ge 0\}$$
 sequence of i.i.d random variables
• $\varphi \equiv$ shift operator on $\Omega = \mathbb{R}^{\mathbb{N}}$
Then φ is ergodic.

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Proof

Measurability of an invariant set: Let $A \in \mathcal{I}$. Then

$$A = \{\omega \in \Omega; \varphi(\omega) \in A\} \implies A \in \mathcal{F}'_1.$$

Tail σ -field: Iterating the previous relation we get

 $A \in \mathcal{I} \implies A \in \mathcal{T}$

Hence by Kolmogorov's 0-1 law we get that ${\mathcal I}$ is trivial

Ergodicity for Markov chains

Example 25.

Let

- $X = \{X_n; n \ge 0\}$ MCH on countable state space S
- Transition probability: p(x, A)
- $\varphi \equiv \text{shift operator on } \Omega = S^{\mathbb{N}}$
- Hypothesis 1: unique stationary distribution π
- Hypothesis 2: $\pi(x) > 0$ for all $x \in S$
- Hypothesis 3: $\mathcal{L}(X_0) = \pi$

Then

- If X is not irreducible, θ is not ergodic
- **2** If X is irreducible, θ is ergodic

Proof

Basic Markov chains facts:

- Since $\pi(x) > 0$ for all $x \in S$, all states are recurrent
- State space decomposition:

 $S = \cup_{j \leq J} R_j$, where R_j disjoint irreducible sets

Non irreducible case: If $J \ge 2$ we have

$$X_0 \in R_j \quad \Longleftrightarrow \quad X_n \in R_j \text{ for all } n \geq 0.$$

Therefore for all $j \leq J$ we get:

- $\mathbf{1}_{(X_0 \in R_j)} = \mathbf{1}_{(X_0 \in R_j)} \, \mathbf{1}_{(X_1 \in R_j)} = \mathbf{1}_{(X_1 \in R_j)}$
- Iterating we get $(X_0 \in R_j) \in \mathcal{I}$
- $\pi(X_0 \in R_j) \in (0,1)$ if $J \geq 2$

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Proof (2) General relation for $\mathbf{1}_A$:

• For $A \in \mathcal{I}$, we have $\mathbf{1}_A = \mathbf{1}_A \circ \theta_n$

• We set
$$\mathcal{F}_n = \sigma(X_0, \ldots, X_n)$$

• We define
$$h(x) = \mathbf{E}_x[\mathbf{1}_A]$$

Then

$$\mathbf{E}_{\pi} \left[\mathbf{1}_{A} | \mathcal{F}_{n} \right] \stackrel{\text{invariance}}{=} \mathbf{E}_{\pi} \left[\mathbf{1}_{A} \circ \theta_{n} | \mathcal{F}_{n} \right] \stackrel{\text{Markov prop}}{=} h(X_{n})$$

Levy's 0-1 law: Let

•
$$\mathcal{F}_n \nearrow \mathcal{F}_\infty$$

•
$$A \in \mathcal{F}_{\infty}$$

Then a.s and in $L^1(\Omega)$ we have

$$\lim_{n\to\infty}\mathbf{E}\left[\mathbf{1}_A|\,\mathcal{F}_n\right]=\mathbf{1}_A$$

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Proof (3)

Limit of $h(X_n)$: Recall

$$\mathbf{E}_{\pi} \left[\mathbf{1}_{A} | \mathcal{F}_{n} \right] = h(X_{n}), \text{ and } \lim_{n \to \infty} \mathbf{E} \left[\mathbf{1}_{A} | \mathcal{F}_{n} \right] = \mathbf{1}_{A}$$

Thus

$$\lim_{n\to\infty}h(X_n)=\mathbf{1}_A.$$
 (3)

Irreducible case: If X irreducible and $\pi(y) > 0$ for all $y \in S$, then

- $h(X_n) = h(y)$ infinitely often for all $y \in S$
- According to (2) we thus have h = Cst

We have thus found that whenever $A \in \mathcal{I}$,

$$\mathbf{1}_{\mathcal{A}}(\omega) = \mathsf{Cst} \in \{0,1\}$$

Ergodicity for rotations of the circle

Example 26.

We consider $\lambda \equiv$ Lebesgue measure and:

- $(\Omega, \mathcal{F}, \mathbf{P}) = ([0, 1], \text{ Borel sets}, \lambda)$
- θ ∈ (0, 1)

•
$$X = \{X_n; n \ge 0\}$$
 with

 $X_n(\omega) = (\omega + n\theta) \mod 1 = \omega + n\theta - [\omega + n\theta].$

• $\varphi \equiv \text{shift transformation}$

Then the following holds true:

- **1** If θ is rational, then φ is not ergodic
- **2** If θ is irrational, then φ is ergodic

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Transformation of an ergodic sequence

Theorem 27.

Let

- $X = \{X_n; n \ge 0\}$ ergodic sequence
- $g: \mathbb{R}^{\mathbb{N}} \to \mathbb{R}$ measurable

We consider a sequence Y defined by

$$Y_k = g\left(\{X_{k+n}; n \ge 0\}\right).$$

Then Y is an ergodic sequence.

Proof

Inverse image: Let $B \in \mathcal{R}^{\mathbb{N}}$. Recall that

$$\left\{\omega \in \mathbb{R}^{\mathbb{N}}; \; Y \in B\right\} = \left\{\omega \in \mathbb{R}^{\mathbb{N}}; \; X \in A
ight\},$$

where

$$A = \left\{ \omega \in \mathbb{R}^{\mathbb{N}}; g(X) \in B
ight\}.$$

Consequence for ergodicity: If B satisfies

$$\left\{\omega \in \mathbb{R}^{\mathbb{N}}; (Y_n)_{n \geq 0} \in B\right\} = \left\{\omega \in \mathbb{R}^{\mathbb{N}}; (Y_{1+n})_{n \geq 0} \in B\right\}$$

Then A satisfies

$$\left\{\omega\in\mathbb{R}^{\mathbb{N}};\,(X_{n})_{n\geq0}\in A
ight\}=\left\{\omega\in\mathbb{R}^{\mathbb{N}};\,(X_{1+n})_{n\geq0}\in A
ight\}$$

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Proof (2)

Conclusion: Since A satisfies

$$\left\{\omega \in \mathbb{R}^{\mathbb{N}}; (X_n)_{n \geq 0} \in A\right\} = \left\{\omega \in \mathbb{R}^{\mathbb{N}}; (X_{1+n})_{n \geq 0} \in A\right\},\$$

we have

$$P(X \in A) \in \{0, 1\}.$$

Therefore:

 $\mathbf{P}(Y \in B) \in \{0,1\}.$

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Ergodicity for Bernoulli shifts



Proof

Representation for Y: Recall that we have defined

• {
$$X_i$$
; $i \ge 1$ } i.i.d with common law $\mathcal{B}(1/2)$
• $g(x) = \sum_{i\ge 1} x_i 2^{-(i+1)}$ defined for $x \in \{0, 1\}^{\mathbb{N}}$
• $g_k(x) = g(\{x_{k+i}; i \ge 1\})$ defined for $x \in \{0, 1\}^{\mathbb{N}}$
• $Y_k = g_k(X)$

Then

$$Y_0 \sim \lambda$$
, and $Y_n(\omega) = 2Y_{n-1} \mod 1$.

Ergodicity of Y: We have

• X ergodic

•
$$Y_k = g_k(X)$$

Hence owing to Theorem 27, Y is a stationary sequence.

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2 Ergodic theorem

3 Recurrence

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Birkhoff's ergodic theorem



Ergodic theorem for ergodic maps



Maximal ergodic lemma



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Proof

Lower bound for X: We will prove that for n = 1, ..., k

$$X(\omega) \ge S_n(\omega) - M_k(\varphi(\omega)) \tag{4}$$

Relation (4) for n = 1: We have $S_1(\omega) = X(\omega)$ and $M_k(\varphi(\omega)) \ge 0$. Thus $X(\omega) \ge S_1(\omega) - M_k(\varphi(\omega))$

Relation (4) for n = 2, ..., k: We have $M_k(\varphi(\omega)) \ge S_j(\varphi(\omega))$ for j = 1, ..., k. Thus

$$X(\omega) + M_k(\varphi(\omega)) \ge X(\omega) + S_j(\varphi(\omega)) = S_{j+1}(\omega)$$

and

$$X(\omega) \geq S_{j+1}(\omega) - M_k(\varphi(\omega))$$

Proof (2)

Consequence of relation (4):

$$X(\omega) \geq \max \{S_1(\omega), \dots, S_k(\omega)\} - M_k(\varphi(\omega))$$

Integration of the previous relation:

$$\mathbf{E} \begin{bmatrix} X \mathbf{1}_{(M_k > 0)} \end{bmatrix} \geq \int_{\{M_k > 0\}} \left[\max \{ S_1(\omega), \dots, S_k(\omega) \} - M_k(\varphi(\omega)) \right] d\mathbf{P}$$

=
$$\int_{\{M_k > 0\}} \left[M_k(\omega) - M_k(\varphi(\omega)) \right] d\mathbf{P}$$

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Proof (3)

Conclusion: We have seen

$$\mathsf{E}\left[X\mathbf{1}_{(M_k>0)}\right] \geq \int_{\{M_k>0\}} \left[M_k(\omega) - M_k(\varphi(\omega))\right] \, d\mathsf{P}$$

In addition on $\{M_k > 0\}^c$ we have

We thus get

$$\mathsf{E}\left[X\mathbf{1}_{(M_k>0)}\right] \geq \int \left[M_k(\omega) - M_k(\varphi(\omega))\right] \, d\mathsf{P} \stackrel{\varphi \text{ invariant}}{=} 0$$

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Proof of Theorem 29

Reduction to $\mathbf{E}[X | \mathcal{I}] = 0$: We

• Set
$$\hat{X} = X - \mathbf{E}[X|\mathcal{I}]$$

• Recall that $\mathbf{E}[X|\mathcal{I}]$ is invariant

Therefore we have

$$\frac{1}{n}\sum_{j=0}^{n}X(\varphi^{j}(\omega)) - \mathbf{E}\left[X|\mathcal{I}\right] = \frac{1}{n}\sum_{j=0}^{n}\left(X(\varphi^{j}(\omega)) - \mathbf{E}\left[X|\mathcal{I}\right]\right)$$
$$= \frac{1}{n}\sum_{j=0}^{n}\left[X - \mathbf{E}\left\{X|\mathcal{I}\right\}\right](\varphi^{j}(\omega)) = \frac{1}{n}\sum_{j=0}^{n}\hat{X}(\varphi^{j}(\omega))$$

We can thus prove Theorem 29 for X such that $\mathbf{E}[X|\mathcal{I}] = 0$

Image: A matrix

Proof of Theorem 29 (2)

Sufficient condition: We define

- $\bar{X} = \limsup_n \frac{S_n}{n}$
- $D = \{\omega; \, \bar{X}(\omega) > \varepsilon\}$ for $\varepsilon > 0$

We wish to prove that

 $\mathbf{P}(D)=0$

Invariance of D: Since \bar{X} is invariant, we have $D \in \mathcal{I}$

Proof of Theorem 29 (3)

Notation: We set

•
$$X^*(\omega) = (X(\omega) - \varepsilon)\mathbf{1}_D(\omega)$$

• $X_j^* = X^*(\varphi^j(\omega))$
• $S_n^* = \sum_{j=0}^{n-1} X_j^*$
• $M_k^* = \max\{0, S_1^*, \dots, S_k^*\}$
• $F_n = \{M_n^* > 0\}$ (increasing sequence
• $F = \bigcup_{n \ge 0} F_n$

Relation between F and D: We have

$${\mathcal F}=\{ ext{There exists } n ext{ s.t } X(arphi^n(\omega)) > arepsilon\} \cap D=D$$

We thus wish to prove that

$$\mathbf{P}(D)=\mathbf{P}(F)=0$$

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Proof of Theorem 29 (4) Proof of $\mathbf{E}[X^* \mathbf{1}_D] \ge 0$: Owing to Lemma 31 we have

 $\mathsf{E}\left[X^*\,\mathbf{1}_{F_n}\right]\geq 0$

By dominated convergence we get:

$$\mathbf{E}\left[X^*\,\mathbf{1}_D\right] = \mathbf{E}\left[X^*\,\mathbf{1}_F\right] \ge 0$$

Proof of $\mathbf{P}(D) = 0$: Since $D \in \mathcal{I}$ and $\mathbf{E}[X|\mathcal{I}] = 0$ we get

$$\mathbf{E}[X^* \mathbf{1}_D] = \mathbf{E}[(X - \varepsilon) \mathbf{1}_D] = \mathbf{E} \{ \mathbf{E}[X | \mathcal{I}] \mathbf{1}_D \} - \varepsilon \mathbf{P}(D) = -\varepsilon \mathbf{P}(D)$$

We have seen that $\mathbf{E}[X^* \mathbf{1}_D] \ge 0$, which yields

 $\mathbf{P}(D) = 0$

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Proof of Theorem 29 (5)

Almost sure limit of $\frac{S_n}{n}$: For $k \ge 1$ we have seen that

$$\mathbf{P}(D_k) \equiv \mathbf{P}\left(\left\{\omega; \limsup_n \frac{S_n}{n} > \frac{1}{k}\right\}\right) = 0$$

Taking limits as $k \to \infty$ we get:

$$\mathbf{P}\left(\left\{\omega;\,\limsup_n\frac{S_n}{n}>0\right\}\right)=0$$

Since the same result is true for the r.v -X we end up with:

$$\mathbf{P}\left(\left\{\omega;\,\limsup_n\frac{S_n}{n}=0\right\}\right)=1$$

Proof of Theorem 29 (6)

Truncation procedure: For the convergence in $L^1(\Omega)$ we set

$$X_M^1 = X \mathbf{1}_{(|X| \le M)}, \quad \text{and} \quad X_M^2 = X \mathbf{1}_{(|X| > M)}$$

and

$$A_n = \frac{1}{n} \sum_{m=0}^{n-1} X(\varphi^m(\omega)) - \mathbf{E}[X|\mathcal{I}]$$

Then for $n \ge 1$ we have

$$\mathsf{E}\left[\left|A_{n}\right|\right] \leq \mathsf{E}\left[\left|A_{n}^{1}\right|\right] + \mathsf{E}\left[\left|A_{n}^{2}\right|\right],$$

where for j = 1, 2 we have defined:

$$A_n^j = \frac{1}{n} \sum_{m=0}^{n-1} X_M^j(\varphi^m(\omega)) - \mathbf{E}[X_M^j | \mathcal{I}]$$

Image: A matrix

Proof of Theorem 29 (7)

Limit for A_n^1 : We have

• a.s $-\lim_{n\to\infty} A_n^1 = 0$ • $|A_n^1| \le 2M$

Therefore by dominated convergence we have:

 $\lim_{n\to\infty} \mathbf{E}\left[\left|A_n^1\right|\right] = 0$

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Proof of Theorem 29 (8)

Limit for A_n^2 : We have

$$\mathbf{E}\left[\left|A_{n}^{2}\right|\right] \leq \frac{1}{n}\sum_{m=0}^{n-1}\mathbf{E}\left[\left|X_{M}^{2}(\varphi^{m}(\omega))\right|\right] + \mathbf{E}\left[\mathbf{E}[\left|X_{M}^{2}\right| \mid \mathcal{I}]\right] \\ \leq 2\mathbf{E}\left[\left|X_{M}^{2}\right|\right]$$

In addition, by dominated convergence we have:

$$\lim_{n\to\infty} \mathbf{E}\left[|X_M^2|\right] = 0$$

Therefore

 $\lim_{n\to\infty}\mathbf{E}\left[\left|A_n^2\right|\right]=0$

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Proof of Theorem 29 (9)

Conclusion for the $L^1(\Omega)$ convergence: We have seen

$$\lim_{n \to \infty} \mathbf{E}\left[\left|A_n^1\right|\right] = 0, \quad \text{and} \quad \lim_{n \to \infty} \mathbf{E}\left[\left|A_n^2\right|\right] = 0.$$

We thus get

$$L^{1}(\Omega) - \lim_{n \to \infty} \frac{1}{n} \sum_{m=0}^{n-1} X_{M}^{j}(\varphi^{m}(\omega)) = \mathbf{E}[X_{M}^{j} | \mathcal{I}]$$

Image: A matrix

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LLN for iid random variables



Remark:

W.r.t the usual LLN, we have obtained the $L^1(\Omega)$ convergence
Previous results: We have seen that

- X is an ergodic sequence
- ② $\mathcal{I} \subset \mathcal{T}$ and \mathcal{T} is trivial

Conclusion: Theorem 29 applies and can be read as

 $\lim_{n\to\infty}\bar{X}_n={\sf E}[X_0|\,{\mathcal I}]={\sf E}[X_0],\quad \text{a.s and in } L^1(\Omega)$

Remark: The $L^1(\Omega)$ convergence can also be obtained as follows

• Result: If $\mathbf{P} - \lim_{n \to \infty} Y_n = Y$, then $\hookrightarrow L^1(\Omega) - \lim_{n \to \infty} Y_n = Y$ iff $\lim_{n \to \infty} \mathbf{E}[|Y_n|] = \mathbf{E}[|Y|]$

• Apply this result successively to $Y_n = X_n^+$ and $Y_n = X_n^-$

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LLN for Markov chains

Example 33.

Let

- $X = \{X_n; n \ge 0\}$ MCH on countable state space S
- Hypothesis 1: unique stat. dist. π and $\mathcal{L}(X_0) = \pi$
- Hypothesis 2: $\pi(x) > 0$ for all $x \in S$
- Hypothesis 3: X irreducible
- Hypothesis 4: $f: S \to \mathbb{R}$ satisfies $f \in L^1(\pi)$

Then

$$\lim_{n\to\infty}\frac{1}{n}\sum_{j=1}^n f(X_j) = \sum_{x\in S} f(x) \pi(x), \quad \text{a.s and in } L^1(\Omega)$$

Previous results: We have seen that

- X is an ergodic sequence
- **2** Therefore f(X) is an ergodic sequence
- ${f 0}$ ${\cal I}$ is trivial whenever X is irreducible

Conclusion: Theorem 29 applies and can be read as

$$\lim_{n\to\infty}\frac{1}{n}\sum_{j=1}^n f(X_j) = \mathbf{E}[f(X_0)|\mathcal{I}] = \mathbf{E}[f(X_0)], \quad \text{a.s and in } L^1(\Omega)$$

Remark: W.r.t the usual LLN for Markov chains, we have obtained the $L^1(\Omega)$ convergence

LLN for rotation of the circle

Example 34.

We consider $\lambda \equiv$ Lebesgue measure and:

- $(\Omega, \mathcal{F}, \mathbf{P}) = ([0, 1], \text{ Borel sets}, \lambda)$
- $heta \in (0,1) \cap \mathbb{Q}^c$
- A Borel subset of [0,1]
- $X = \{X_n; n \ge 0\}$ with

$$X_n(\omega) = (\omega + n\theta) \mod 1 = \omega + n\theta - [\omega + n\theta].$$

Then

$$\lim_{n\to\infty}\frac{1}{n}\sum_{j=1}^n \mathbf{1}_A(X_n) = |A|, \quad \text{a.s and in } L^1(\Omega)$$

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Deterministic LLN for rotation of the circle



Proof:

- Start from Example 34
- Additional ingredient based on density arguments

Benford's law

 Proposition 36.

 For $k \in \{1, ..., 9\}$ we have

 $\lim_{n \to \infty} \frac{1}{n} \sum_{m=1}^{n} \mathbf{1}_{(1^{\text{st digit of } 2^m = k)} = \log_{10} \left(\frac{k+1}{k}\right)$



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Notation: We set

Expression 1 for $\log_{10}(2^m)$: we have

 $\log_{10}(2^m) = m\theta$

Expression 2 for $\log_{10}(2^m)$: First digit of 2^m is k_0 iff

 $2^m = k_0 10^{lpha} + k_1$, with $lpha \ge 0$, $k_1 < 10^{lpha}$

and

$$\log_{10}(2^m) = \alpha + \log_{10}\left(k_0 + \frac{k_1}{10^{\alpha}}\right)$$

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Proof (2)

Ergodic sequence: We have seen that first digit of 2^m is k_0 iff

$$m \theta = \alpha + \log_{10}\left(k_0 + \frac{k_1}{10^{lpha}}\right),$$

which is equivalent to

 $m \theta \mod 1 \in A_k$

Thus

$$\lim_{n\to\infty}\frac{1}{n}\sum_{m=1}^{n}\mathbf{1}_{(1^{\text{st}}\text{ digit of }2^m=k)}=\lim_{n\to\infty}\frac{1}{n}\sum_{m=1}^{n}\mathbf{1}_{(m\theta \mod 1\in A_k)}$$

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Proof (3)

Conclusion: We have

$$\theta = \log_{10}(2) \in \mathbb{Q}^{c}$$

We can thus apply Theorem 35 with x = 0 and $A = A_k$. We get

$$\lim_{n\to\infty}\frac{1}{n}\sum_{m=1}^{n}\mathbf{1}_{(m\theta \mod 1\in A_k)} = |A_k| = \log_{10}\left(\frac{k+1}{k}\right)$$

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Image: A matrix

Outline

Definitions and examples

- Preliminaries on Markov chains
- Examples of stationary sequences
- Notion of ergodicity

2 Ergodic theorem



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Number of points visited

Theorem 37.
Let
• X stationary sequence with values in
$$\mathbb{R}^d$$

• $S_n = \sum_{j=0}^{n-1} X_j$
• $A = \{S_k \neq 0 \text{ for all } k \ge 1\}$
• $R_n = |\{S_1, \dots, S_n\}| \equiv \operatorname{Card}(\{S_1, \dots, S_n\})$
Then
 $\lim_{n \to \infty} \frac{R_n}{n} = \mathbf{E}[\mathbf{1}_A | \mathcal{I}], \text{ almost surely}$

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Lower bound: We have

$$S_j(\varphi^m(\omega)) = S_{j+m}(\omega) - S_m(\omega)$$

Thus

$$S_j(\varphi^m(\omega)) \neq 0 \quad \Longleftrightarrow \quad S_{j+m}(\omega) \neq S_m(\omega)$$

and

$$\mathbf{1}_{A}(\varphi^{m}(\omega)) = \{S_{j+m}(\omega) \neq S_{m}(\omega) \text{ for all } j \geq 1\}$$

We have thus obtained

$$R_n \geq \sum_{m=1}^n \mathbf{1}_A(\varphi^m(\omega)).$$

and owing to the ergodic theorem we get

$$\liminf_{n\to\infty}\frac{R_n}{n}\geq\liminf_{n\to\infty}\frac{1}{n}\sum_{m=1}^n\mathbf{1}_A(\varphi^m(\omega))=\mathbf{E}\left[\mathbf{1}_A|\mathcal{I}\right]$$

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Proof (2) Upper bound: For $k \ge 1$ set

$$A_k = \{S_j \neq 0 \text{ for all } 1 \leq j \leq k\}.$$

Then

$$\mathbf{1}_{A_k}(\varphi^m(\omega)) = \{S_j(\omega) \neq S_m(\omega) \text{ for all } m+1 \leq j \leq m+k\}$$

and

$$R_n \leq k + \sum_{m=1}^{n-k} \mathbf{1}_{A_k}(\varphi^m(\omega)).$$

Applying the ergodic theorem again we get

$$\liminf_{n\to\infty}\frac{R_n}{n}\leq\limsup_{n\to\infty}\frac{1}{n}\sum_{m=1}^{n-k}\mathbf{1}_{A_k}(\varphi^m(\omega))=\mathbf{E}\left[\mathbf{1}_{A_k}|\mathcal{I}\right]$$
(5)

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Proof (3)

Limit in k: We have $k \mapsto A_k$ decreasing and

$$\bigcap_{k=1}^{\infty} A_k = A.$$

Therefore by monotone convergence we get

$$\lim_{k\to\infty} \mathsf{E}\left[\mathbf{1}_{\mathcal{A}_k} | \mathcal{I}\right] = \mathsf{E}\left[\mathbf{1}_{\mathcal{A}} | \mathcal{I}\right].$$

Thanks to (5) we have thus proved that

$$\limsup_{n\to\infty}\frac{R_n}{n}\leq \mathsf{E}\left[\mathbf{1}_{A}|\mathcal{I}\right]$$

Recurrence criterion in $\ensuremath{\mathbb{Z}}$



Standing assumption: $\mathbf{E}[X_1 | \mathcal{I}] = 0$, which yields

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$$\lim_{n \to \infty} \frac{S_n}{n} = 0.$$
 (6)

Relation for max S: For all K > 1 we have

$$\limsup_{n \to \infty} \max_{1 \le k \le n} \frac{|S_k|}{n} = \limsup_{n \to \infty} \max_{K \le k \le n} \frac{|S_k|}{n} \le \max_{k \ge K} \frac{|S_k|}{k}$$

Taking $K \to \infty$ and thanks to (6) we get

 $n \rightarrow \infty 1 \le k \le n$ *n*

$$\lim_{k \to \infty} \max \frac{|S_k|}{|S_k|} = 0$$

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(7)

Proof (2) Consequence on R_n : In d = 1 we have

$$R_n = \mathsf{Range}\left\{X_j; \ 0 \leq j \leq n
ight\}.$$

Hence

$$R_n \leq 1+2 \max_{1 \leq k \leq n} |S_k|,$$

and according to (7) we get

$$\lim_{n \to \infty} \max_{1 \le k \le n} \frac{|R_n|}{n} \stackrel{\text{Theorem 37}}{=} \mathbf{E} \left[\mathbf{1}_A | \mathcal{I} \right] \stackrel{\text{Relation (7)}}{=} 0$$

Conclusion for item 1:

We take expectations in previous relation, which yields

$$\mathbf{P}(A) = 0$$

Proof (3)

Notation: We set

$$\begin{array}{rcl} F_{j} & = & \{S_{i} \neq 0 \mbox{ for } i < j \mbox{ and } S_{j} = 0\} \\ G_{j,k} & = & \{S_{j+i} - S_{j} \neq 0 \mbox{ for } i < k \mbox{ and } S_{j+k} - S_{j} = 0\} \end{array}$$

Simple relations on F_j , $G_{j,k}$:

- Since $\mathbf{P}(A) = 0$ we have $\sum_{k \ge 1} \mathbf{P}(F_k) = 1$
- 3 By stationarity, $\mathbf{P}(G_{j,k}) = \mathbf{P}(F_k)$
- Solution Hence we also have $\sum_{k\geq 1} \mathbf{P}(G_{j,k}) = 1$
- For fixed j, the sets $G_{j,k}$ are disjoint

Proof (4)

More relations on F_j , $G_{j,k}$: We have obtained

 $\sum_{k\geq 1} \mathsf{P}\left(F_{j} \cap G_{j,k}\right) = \mathsf{P}\left(F_{j}\right), \text{ and } \sum_{j,k\geq 1} \mathsf{P}\left(F_{j} \cap G_{j,k}\right) = 1$

Conclusion on recurrence: We have

$$\bigcup_{j,k\geq 1}F_j\cap G_{j,k}=\{S_n=0 \text{ at least two times}\}$$

Therefore

$$\mathbf{P}(S_n = 0 \text{ at least two times}) = 1$$

Generalization:

- With sets $G_{j_1,...,j_m}$
- Taking $m \to \infty$

Extension by Kac

Theorem 39.

Let

• X stationary sequence with values in (S, S)• $A \in S$ and $S_n = \sum_{i=0}^{n-1} X_i$ • $T_0 = 0$ and $T_n = \inf\{m > T_{n-1}; X_m \in A\}$ • $t_n = T_n - T_{n-1}$ • Hypothesis: $\mathbf{P}(X_n \in A \text{ at least once}) = 1$ Then the following holds true: **1** Under $\mathbf{P}(\cdot | X_0 \in A)$ the sequence (t_n) is stationary 2 We have $\mathbf{E}[T_1|X_0 \in A] = \frac{1}{\mathbf{P}(X_0 \in A)}$

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