

Summary on Poisson process . $N(t)$ is PP if

(i) $N(t) \in \mathbb{N}$, $N(0) = 0$

(ii) $t \mapsto N(t)$ is \nearrow

(iii) $\mathbb{P}(N(t+h) = n+m \mid N(t) = n)$

$$= \begin{cases} 1 - \lambda h + g_1(h) & \text{if } m = 0 \\ \lambda h + g_2(h) & \text{if } m = 1 \\ g_3(h) & \text{if } m > 1 \end{cases}$$

(iv) $N(t) - N(s) \perp\!\!\!\perp$ "past"

with $\lim_{h \rightarrow 0} \frac{|g_1(h)| + |g_2(h)| + |g_3(h)|}{h} \rightarrow 0$

Notation: we say $g_1(h), g_2(h), g_3(h) = o(h)$

First important result. For all $t \geq 0$

$$N(t) \sim P(\lambda t)$$

Intuition This stems from

$$\text{Bin}(n, \frac{\mu}{n}) \xrightarrow{n \rightarrow \infty} P(\mu)$$

Proof of $\text{Bin}(n, \frac{\mu}{n}) \longrightarrow \mathcal{P}(\mu)$ means convergence
in distribution

It boils down to prove, for $k \in \mathbb{N}$, that

$$\mathbb{P}(X_n = k) \longrightarrow \mathbb{P}(X = k)$$

where $X_n \sim \text{Bin}(n, \frac{\mu}{n})$, $X \sim \mathcal{P}(\mu)$

Computation

$$\mathbb{P}(X_n = k) = \binom{n}{k} \left(\frac{\mu}{n}\right)^k \left(1 - \frac{\mu}{n}\right)^{n-k}$$

$$= \frac{n!}{k! (n-k)!} \frac{\mu^k}{n^k} \left(1 - \frac{\mu}{n}\right)^{n-k}$$

$$= \frac{\mu^k}{k!} \frac{n!}{(n-k)! n^k} \left(1 - \frac{\mu}{n}\right)^n \left(1 - \frac{\mu}{n}\right)^{-k}$$

$$P(X_n = k) = \frac{e^{\mu k}}{k!} \underbrace{\frac{n!}{(n-k)! n^k}}_{A(n)} \underbrace{\left(1 - \frac{\mu}{n}\right)^n}_{B(n)} \underbrace{\left(1 - \frac{\mu}{n}\right)^{-k}}_{C(n)}$$

Then

$$A(n) = \frac{n(n-1)\dots(n-k+1)}{n^k} = \frac{n^k + \overbrace{P_{k-1}(n)}^{\text{polynomial degree } k-1}}{n^k}$$

$$B(n) = e^{n \ln\left(1 - \frac{\mu}{n}\right)} \xrightarrow{n \rightarrow \infty} e^{-\mu}$$

$$C(n) = \frac{1}{\left(1 - \frac{\mu}{n}\right)^k} \xrightarrow{n \rightarrow \infty} 1 \quad (\text{cont. function})$$

conclusion

$$P(X_n = k) \longrightarrow \frac{\mu^k}{k!} e^{-\mu} = P(X = k)$$

Proof of Thm 2

no more than 1 arrival in
small time

$$N(t+h) \geq N(t)$$

$$P(N(t+h) = j)$$

$$= \sum_{i=0}^j P(N(t+h) = j \mid N(t) = i) P(N(t) = i)$$

h small

$$\approx P(N(t+h) = j \mid N(t) = j) P(N(t) = j)$$

$$+ P(N(t+h) = j \mid N(t) = j-1) P(N(t) = j-1) + o(h)$$

$$= (1 - \lambda h + o(h)) P(N(t) = j)$$

$$+ (\lambda h + o(h)) P(N(t) = j-1) + o(h)$$

Notation we let $p_j(t) = P(N(t) = j)$.

we get

$$p_j(t+h) = (1 - \lambda h) p_j(t) + \lambda h p_{j-1}(t) + o(h)$$

$$P_j(t) = \mathbb{P}(N(t) = j)$$

We have found

$$P_j(t+h) = (1-\lambda h) P_j(t) + \lambda h P_{j-1}(t) + o(h)$$

$$\Leftrightarrow P_j(t+h) - P_j(t) = -\lambda h P_j(t) + \lambda h P_{j-1}(t) + o(h)$$

$$\Leftrightarrow \frac{1}{h} (P_j(t+h) - P_j(t)) = -\lambda P_j(t) + \lambda P_{j-1}(t) + o(1)$$

Take $h \rightarrow 0$. we get

$$P'_j(t) = -\lambda P_j(t) + \lambda P_{j-1}(t)$$

we get an ∞ (in j) system of **linear** differential equations

2 strategies : solve by induction

Generating functions

$$= P(N(0)=0)$$

system

$$p'_j(t) = -\lambda p_j(t) + \lambda p_{j-1}(t)$$

Case $j=0$ Since $p_{-1}(t)=0$, the eq is

$$p'_0(t) = -\lambda p_0(t), \quad p_0(0) = 1$$

$$\Leftrightarrow \frac{p'_0(t)}{p_0(t)} = -\lambda \quad (\text{reparable})$$

$$\Rightarrow \ln(p_0(t)) = -\lambda t + C_1$$

$$p_0(t) = C_2 e^{-\lambda t}$$

with $p_0(0)=1$, we get $X \sim P(\lambda t)$

$$p_0(t) = e^{-\lambda t} = P(\bar{X}=0)$$

$$P'_j(t) = -\lambda P_j(t) + \lambda P_{j-1}(t)$$

Diff für $G_t(s) = \sum_{j=0}^{\infty} P_j(t) s^j$. we have

$$\begin{aligned} \partial_t G_t(s) &= P'_0(t) + \sum_{j=1}^{\infty} P'_j(t) s^j \\ &= -\lambda P_0(t) + \sum_{j=1}^{\infty} (-\lambda P_j(t) + \lambda P_{j-1}(t)) s^j \\ &= -\lambda P_0(t) + \sum_{j=1}^{\infty} -\lambda P_j(t) s^j \\ &\quad + s \sum_{j=1}^{\infty} \lambda P_{j-1}(t) s^{j-1} \\ &= -\lambda G_t(s) + \lambda s G_t(s) \end{aligned}$$

$$\Rightarrow \partial_t G_t(s) = \lambda(s-1) G_t(s), \quad G_0(s) = 1$$

$$\Rightarrow G_t(s) = e^{\lambda(s-1)t} \Rightarrow P_j(t) \in \mathcal{P}(\lambda t)$$

Proof of Theorem 2 (1)

Conditioning on a small interval: We have

$$\begin{aligned} & \mathbf{P}(N(t+h) = j) \\ &= \sum_{i \in S} \mathbf{P}(N(t+h) = j | N(t) = i) \mathbf{P}(N(t) = i) \\ &= \sum_{i \in S} \mathbf{P}((j-i) \text{ arrivals in } (t, t+h]) \mathbf{P}(N(t) = i) \\ &= \mathbf{P}(\text{no arrivals in } (t, t+h]) \mathbf{P}(N(t) = j) \\ &\quad + \mathbf{P}(\text{one arrival in } (t, t+h]) \mathbf{P}(N(t) = j-1) + o(h) \\ &= (1 - \lambda h) \mathbf{P}(N(t) = j) + \lambda h \mathbf{P}(N(t) = j-1) + o(h) \end{aligned}$$

Proof of Theorem 2 (2)

Probability as a function: We set

$$p_j(t) = \mathbf{P}(N(t) = j)$$

Equation on small intervals: We have seen

$$p_0(t+h) = (1 - \lambda h) p_0(t) + o(h)$$

$$p_j(t+h) = \lambda h p_{j-1}(t) + (1 - \lambda h) p_j(t) + o(h)$$

Equivalent form with differences:

$$p_0(t+h) - p_0(t) = -\lambda h p_0(t) + o(h)$$

$$p_j(t+h) - p_j(t) = \lambda h (p_{j-1}(t) - p_j(t)) + o(h)$$

Proof of Theorem 2 (3)

Recall:

$$\begin{aligned}p_0(t+h) - p_0(t) &= -\lambda h p_0(t) + o(h) \\p_j(t+h) - p_j(t) &= \lambda h (p_{j-1}(t) - p_j(t)) + o(h)\end{aligned}$$

Differentiating: We end up with a system of ode's

$$\begin{aligned}p_0'(t) &= -\lambda p_0(t) \\p_j'(t) &= \lambda p_{j-1}(t) - \lambda p_j(t)\end{aligned}$$

Initial condition:

$$p_j(0) = \delta_{j0} \equiv \mathbf{1}_{(j=0)}$$

Proof of Theorem 2 (4)

Recall: We have obtained a system of ode's

$$\begin{aligned}p_0'(t) &= -\lambda p_0(t) \\p_j'(t+h) &= \lambda p_{j-1}(t) - \lambda p_j(t)\end{aligned}$$

A family of generating functions: We set

$$G_t(s) = \mathbf{E} [s^{N(t)}] = \sum_{j=0}^{\infty} p_j(t) s^j$$

Strategy: From the system of ode's

\hookrightarrow deduce a single ode for $t \mapsto G_t(s)$

Proof of Theorem 2 (5)

Differential equation for G : We have

$$\begin{aligned}\frac{\partial G_t(s)}{\partial t} &= \sum_{j=0}^{\infty} p_j'(t) s^j \\ &= -\lambda p_0(t) + \sum_{j=1}^{\infty} (\lambda p_{j-1}(t) - \lambda p_j(t)) s^j \\ &= -\lambda G_t(s) + \lambda s \sum_{j=1}^{\infty} p_{j-1}(t) s^{j-1} \\ &= -\lambda G_t(s) + \lambda s G_t(s) \\ &= \lambda(s-1)G_t(s)\end{aligned}$$

Proof of Theorem 2 (6)

Recall: $u_t \equiv G_t(s)$ verifies

$$u' = \lambda(s - 1)u, \quad u_0 = 1$$

Expression for $G_t(s)$: We find

$$G_t(s) = \exp(\lambda(s - 1)t)$$

Conclusion:

$$N(t) \sim \mathcal{P}(\lambda t),$$

Relation with binomial random variables

Another way to prove $N(t) \sim \mathcal{P}(\lambda t)$:

- 1 Partition $[0, t]$ in subintervals $[(\ell - 1)h, \ell h]$
- 2 On each subinterval, set $Z_\ell = \mathbf{1}_{(\text{arrival in } [(\ell-1)h, \ell h])}$
- 3 We have that $\{Z_\ell; \ell \geq 1\}$ is i.i.d with common law $\mathcal{B}(\lambda h)$
- 4 We have $N(t) \simeq \sum_{\ell=1}^{t/h} Z_\ell$, thus

$$N(t) \simeq \text{Bin} \left(\frac{t}{h}; \lambda h \right) \xrightarrow{h \rightarrow 0} \mathcal{P}(\lambda t)$$