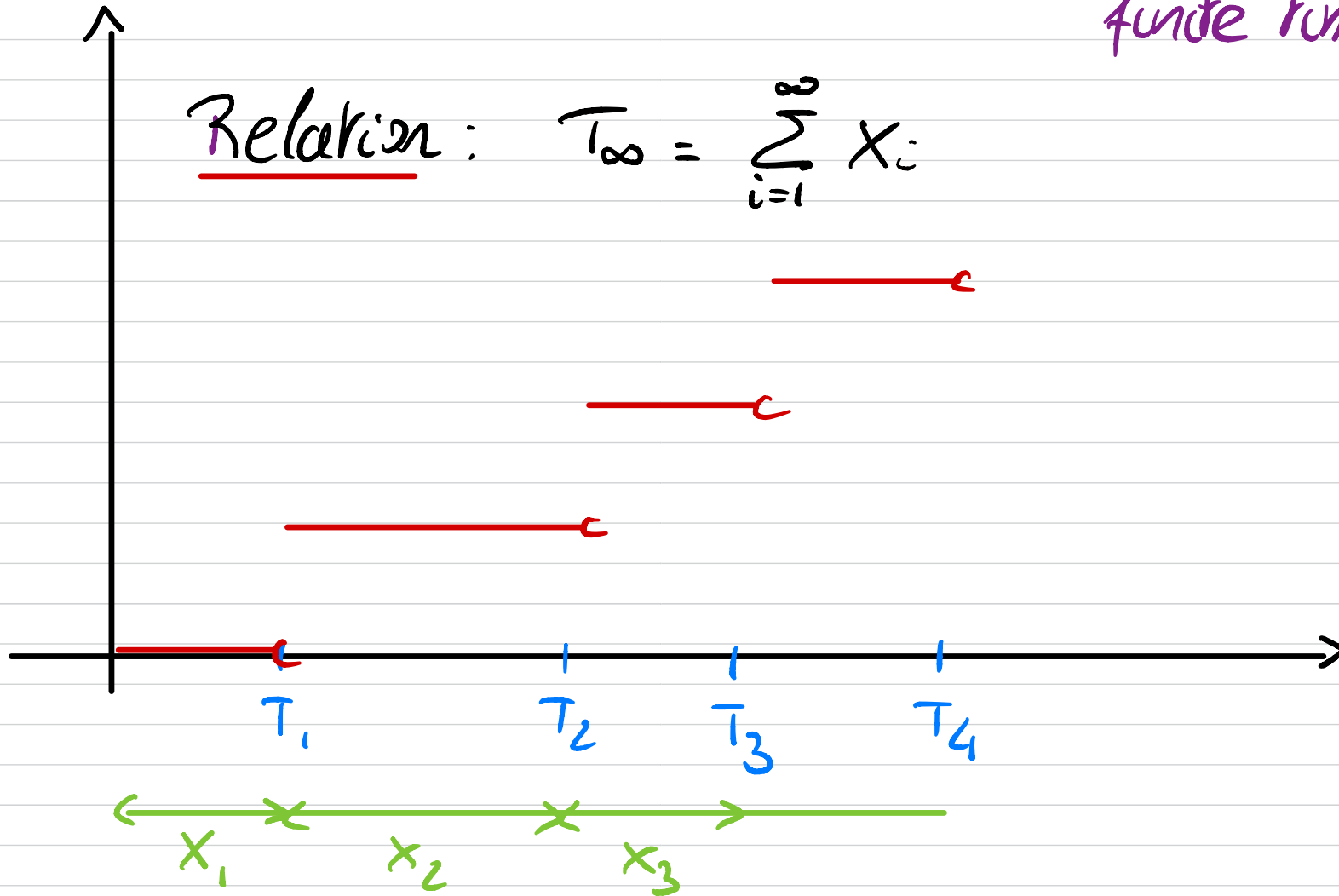


Question: Is  $T_\infty < \infty$ ? (Explosion in finite time)

Relation:  $T_\infty = \sum_{i=1}^{\infty} X_i$



We know that  $X_i$ 's are i.i.d.,  $X_i \sim E(\lambda_{i-1})$

# Sum of exponential random variables

## Proposition 15.

Let

- $\{X_n; n \geq 1\}$  sequence of independent random variables
- Each  $X_n$  is such that  $X_n \sim \mathcal{E}(\lambda_{n-1})$
- $T_\infty = \sum_{n=1}^{\infty} X_n$

Then

$$\mathbf{P}(T_\infty < \infty) = \begin{cases} 0, & \text{if } \sum_{n=1}^{\infty} \frac{1}{\lambda_n} = \infty \\ 1 & \text{if } \sum_{n=1}^{\infty} \frac{1}{\lambda_n} < \infty \end{cases}$$

# Proof of Proposition 15 (1)

Case  $\sum_{n \geq 1} \lambda_n^{-1} < \infty$ : Using Fubini-Tonelli we have

$$\mathbf{E}[T_\infty] = \mathbf{E}\left[\sum_{n=1}^{\infty} X_n\right] = \sum_{n=1}^{\infty} \frac{1}{\lambda_{n-1}} < \infty$$

Thus

$$\mathbf{P}(T_\infty < \infty) = 0$$

Case  $\sum_{n=1}^{\infty} \frac{1}{\lambda_n} = \infty$  . In that case, we

consider a kind of Laplace transform:

$$\mathbb{E}[e^{-T_{\infty}}]$$

$e^{-T_{\infty}}$  is a  $\geq 0$  r.v. If  $\mathbb{E}[e^{-T_{\infty}}] = 0$ ,  
then

$$\mathbb{P}(e^{-T_{\infty}} = 0) = 1$$

or  $\mathbb{E}[e^{-T_{\infty}}] = 0 \Rightarrow e^{-T_{\infty}} = 0$  a.s

$$\mathbb{P}(T_{\infty} = \infty) = 1$$

Reduction: It is enough to prove

$$\mathbb{E}[e^{-T_{\infty}}] = 0$$

Computation for  $E[e^{-T_{\infty}}]$

$$E[e^{-T_{\infty}}] = E\left[\exp\left(-\sum_{n=1}^{\infty} X_n\right)\right]$$
$$= E\left[\prod_{n=1}^{\infty} e^{-X_n}\right] \stackrel{?}{=} \prod_{n=1}^{\infty} E[e^{-X_n}]$$

$$= E\left[\lim_{N \rightarrow \infty} \prod_{n=1}^N e^{-X_n}\right]$$

dominated / monotone convergence

$$= \lim_{N \rightarrow \infty} E\left[\prod_{n=1}^N e^{-X_n}\right]$$

$$= \lim_{N \rightarrow \infty} \prod_{n=1}^N E[e^{-X_n}]$$

Next

$$E[e^{-X_n}] = \int_0^{\infty} \lambda_{n-1} e^{-\lambda_{n-1} t} e^{-t} dt$$

density  $E(\lambda_{n-1})$

$$= \lambda_{n-1} \int_0^{\infty} e^{-(\lambda_{n-1} + 1)t} dt = \frac{\lambda_{n-1}}{\lambda_{n-1} + 1} = \frac{1}{1 + \frac{1}{\lambda_{n-1}}}$$

## Summary

$$E[e^{-T_0}] = \lim_{N \rightarrow \infty} \prod_{n=1}^N \frac{1}{1 + \frac{1}{\alpha_{n-1}}}$$

$$= \prod_{n=1}^{\infty} \frac{1}{1 + \frac{1}{\alpha_{n-1}}}$$

$$= \frac{1}{\prod_{n=1}^{\infty} \left(1 + \frac{1}{\alpha_{n-1}}\right)}$$

## 2nd reduction

$$E[e^{-T_0}] = 0 \Leftrightarrow \prod_{n=1}^{\infty} \left(1 + \frac{1}{\alpha_{n-1}}\right) = \infty \quad \begin{matrix} \equiv U_n \\ \text{circled} \end{matrix}$$

Claim If  $U_n \geq 0$ , then

$$\prod_{n=1}^{\infty} (1 + U_n) = \infty \Leftrightarrow \sum U_n = \infty$$

"Proof" of last claim

$$\ln\left(\prod_{n=1}^{\infty} (1+u_n)\right) = \sum_{n=1}^{\infty} \ln(1+\underbrace{u_n})$$

hyp:  $u_n$  small

$$\approx \sum_{n=1}^{\infty} u_n$$

$$\left(\frac{1}{2}u \leq \ln(1+u) \leq u \text{ if } u \text{ small}\right)$$

# Proof of Proposition 15 (2)

Case  $\sum_{n \geq 1} \lambda_n^{-1} = \infty$ , strategy: We have

$$\begin{aligned} \mathbf{E} [e^{-T_\infty}] = 0 &\implies \mathbf{P} (e^{-T_\infty} = 0) = 1 \\ &\implies \mathbf{P} (T_\infty = \infty) = 1 \end{aligned}$$

We will thus prove

$$\mathbf{E} [e^{-T_\infty}] = 0$$



# Proof of Proposition 15 (3)

Case  $\sum_{n \geq 1} \lambda_n^{-1} = \infty$ , computation: We have

$$\begin{aligned} \mathbf{E} \left[ e^{-T_\infty} \right] &= \mathbf{E} \left[ \prod_{n=1}^{\infty} e^{-X_n} \right] \\ &= \lim_{N \rightarrow \infty} \mathbf{E} \left[ \prod_{n=1}^N e^{-X_n} \right] \quad (\text{monotone convergence}) \\ &= \lim_{N \rightarrow \infty} \prod_{n=1}^N \mathbf{E} \left[ e^{-X_n} \right] \\ &= \lim_{N \rightarrow \infty} \prod_{n=1}^N \frac{1}{1 + \lambda_{n-1}^{-1}} = \left( \prod_{n=1}^{\infty} \left( 1 + \frac{1}{\lambda_{n-1}} \right) \right)^{-1} \end{aligned}$$

# Proof of Proposition 15 (4)

Infinite products: If  $u_n \geq 0$ , then

$$\prod_{n=1}^{\infty} (1 + u_n) = \infty \iff \sum_{n=1}^{\infty} u_n = \infty \quad (3)$$

Pseudo-proof of (3): We have

$$\begin{aligned} \ln \left( \prod_{n=1}^{\infty} (1 + u_n) \right) &= \sum_{n=1}^{\infty} \ln(1 + u_n) \\ &\asymp \sum_{n=1}^{\infty} u_n \end{aligned}$$

# Proof of Proposition 15 (5)

Recall: We have seen

$$\mathbf{E} [e^{-T_\infty}] = \left( \prod_{n=1}^{\infty} \left( 1 + \frac{1}{\lambda_{n-1}} \right) \right)^{-1}$$

Application of (3):

$$\mathbf{E} [e^{-T_\infty}] < \infty \iff \prod_{n=1}^{\infty} \left( 1 + \frac{1}{\lambda_{n-1}} \right) = \infty \iff \sum_{n \geq 1} \lambda_n^{-1} = \infty$$

Conclusion:

$$T_\infty = \infty \iff \sum_{n \geq 1} \lambda_n^{-1} = \infty$$

# Application to birth process

## Proposition 16.

Let

- $N$  birth process
- Intensities  $\{\lambda_j; j \geq -1\}$ , with  $\lambda_{-1} = 0$
- $\{T_n; n \geq 1\}$  arrival times

Then  $N$  is **honest** iff

$$\sum_{n=1}^{\infty} \frac{1}{\lambda_n} = \infty$$

# Final remarks

## Notes before next section:

- 1 Poisson and birth processes are Markov processes  
↪ Due to  $(N(t) - N(s)) \perp\!\!\!\perp \text{Past}$ , given  $N(s) = i$
- 2 They are in fact strong Markov processes  
↪ Definition to be seen later
- 3 Problems can occur due to explosions  
↪ This could not be observed in discrete time

# Outline

- 1 Birth processes and the Poisson process
  - Poisson process
  - Birth processes
- 2 Continuous time Markov chain
  - General definitions and transitions
  - Generators
  - Classification of states

# Outline

- 1 Birth processes and the Poisson process
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- 2 Continuous time Markov chain
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# Vocabulary

## Stochastic process:

- Family  $\{X(t); t \in [0, \infty)\}$  of random variables
- Family evolving in a random but prescribed manner
- Here  $X(t) \in S$ , where  $S$  countable state space with  $N = |S|$

## Markov evolution:

Conditioned on  $X(t)$ ,  
the evolution does not depend on the past



# Markov chain

## Definition 17.

Let

- $X = \{X(t); t \geq 0\}$  stochastic process

We say that  $X$  is a continuous time Markov chain if

$$\begin{aligned} & \mathbf{P}(X(t_n) = j | X(t_1) = i_1, \dots, X(t_{n-1}) = i_{n-1}) \\ &= \mathbf{P}(X(t_n) = j | X(t_{n-1}) = i_{n-1}), \end{aligned}$$

*set of  $> 0$  probab* (above the second part of the equation)  
*discrete r.v.* (below the second part of the equation)

for all

- $0 \leq t_1 < \dots < t_n < \infty$
- $i_1, \dots, i_n, j \in S$

# Differences with discrete time

## Main difference:

- No time unit
- Therefore no exact analogue of  $P$

## Method 1:

- Use infinitesimal calculus
- This leads to infinitesimal generator

## Method 2:

- Embedded chain  $\{X_{T_n}; T_n \text{ arrival times}\}$  is usually a discrete Markov chain

# Birth process as Markov process

## Proposition 18.

Let

- $N$  birth process
- Intensities  $\{\lambda_j; j \geq -1\}$ , with  $\lambda_{-1} = 0$

Then

$N$  is a Markov process